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Is cognitive behavioral therapy more

binge spectrum disorders? A systematic

effective than pharmacotherapy for

review and meta-analysis

Abstract

Objectives: Binge spectrum disorders are prevalent worldwide. Psychiatric and medical comorbidities are common, and societal costs are significant. Evidence-based treatment remains underutilized. Cognitive behavioral therapy is the recommended first-line treatment, but pharmacotherapy may be easier to access.

Interventions: Meta-analytic evidence directly comparing cognitive behavioral therapy with pharmacotherapy is lacking. We aimed to compare the effects of cognitive behavioral therapy interventions with any pharmacological treatment for binge spectrum disorders. We searched PubMed, Embase, CENTRAL, ClinicalTrials.gov and reference lists for randomized controlled trials comparing cognitive behavioral therapy with any pharmacotherapy for bulimia nervosa/binge eating disorder and performed pairwise meta-analytic evaluations.

Primary Outcomes: Primary outcomes are remission and frequency of binges. Secondary outcomes are frequency of purges, response, eating disorder psychopathology, weight/body mass index, depression, anxiety, quality of life and dropouts.

Results: Eleven randomized controlled trials comparing cognitive behavioral therapy with fluoxetine/imipramine/desipramine/methylphenidate/sibutramine were identified (N=531). Cognitive behavioral therapy was superior to antidepressants in terms of remission, frequency of binges and eating disorder psychopathology. There were no statistically significant differences for any of the individual cognitive behavioral therapy vs drug comparisons in terms of response/ depression/anxiety/weight/quality of life/dropouts. Cognitive behavioral therapy was not superior to sibutramine/methylphenidate for the primary outcomes.

Conclusions: Data are scarce, comparisons underpowered and, considering the inherent methodological limitations of psychotherapy trials, questions arise regarding the presumed superiority of cognitive behavioral therapy. Further research is needed.

Keywords

Eating disorders, psychopharmacotherapy, CBT, drug therapy, meta-analysis

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Introduction

Bulimia nervosa (BN) and binge eating disorder (BED), commonly referred to as binge spectrum disorders (BSDs) (Treasure et al., 2020), are among the most prevalent eating disorders worldwide (Erskine and Whiteford, 2018; Van Eeden et al., 2021). BN is characterized by recurrent episodes of binge eating followed by unhelpful compensatory behaviors. These are collectively referred to as purging, including abuse of laxatives, diuretics, enemas and selfinduced vomiting among others, and are aimed at controlling weight (Jain and Yilanli, 2021). BED is also characterized by frequent and recurrent episodes of binge eating, but without a regular pattern of compensatory behaviors (Igbal and Rehman, 2021). BSDs are associated with a high risk of psychiatric and medical comorbidity (Udo and Grilo, 2019), but also functional decline, as well as increased risk of suicide and parasuicidal self-harm behavior (Perkins and Brausch, 2019; Udo et al., 2019). This often has a devastating impact on the individual and their families and is also associated with significant societal costs (Tannous et al., 2021).

Cognitive behavioral therapy (CBT) interventions are considered the first-line treatment for BSDs (Agras, 2019; Treasure et al., 2020), and they are endorsed by major guidelines (Crone et al., 2023; NICE Guideline, 2017). The theoretical basis of CBT for eating disorders relies on clinical observations and social learning theory. This framework suggests that low self-esteem and persistent negative emotions lead to dissatisfaction with body weight and shape, resulting in restricted eating, loss of control over eating and binge eating with or without compensatory behaviors (Agras and Bohon, 2021).

Interpersonal issues and life stressors often trigger binge-eating episodes. Several models and forms of delivery are available (Mulkens and Waller, 2021). CBT interventions are supported by a larger body of evidence compared to any other modality of psychotherapy and are widely used in clinical practice (Peat et al., 2017; Svaldi et al., 2019). However, access to trained therapists (Kazdin et al., 2017), the willingness of patients to engage in therapy (Thompson and Park, 2016), financial and geographical access problems and long-waiting lists (Liu et al., 2022; Thompson and Park, 2016) are important barriers to psychological treatment.

Despite CBT's historical promotion as the 'gold standard', a growing body of literature is critical of this perspective. This literature argues that claims about CBT's superiority lack thorough assessment and fail to acknowledge potential harms or iatrogenic effects (Castro Batic and Hayes, 2020; Hayes and Za'ba, 2021; Lilienfeld, 2007; Parry et al., 2016). In addition, there is increasing evidence that CBT may not be the most suitable approach and can even be potentially harmful for certain marginalized communities, such as autistic individuals (Babb et al., 2022), who are overrepresented in eating disorder populations (Huke et al., 2013).

Access to pharmacotherapy is often easier and more practical. However, medication remains a second-line option for BED (Crone et al., 2023), and only fluoxetine is Food and Drug Administration (FDA)-approved for the treatment of BN, and lisdexamfetamine for BED. Fluoxetine is recommended either as initial treatment or as an augmentation to psychotherapy when the latter alone fails to deliver the desirable clinical outcomes (Crone et al., 2023). Nevertheless, the results of recent meta-analyses indicated that fluoxetine has small clinical effects compared to other treatment options and placebo (Argyrou et al., 2023; Slade et al., 2018). It is also worth mentioning that there is an ongoing and robust debate surrounding the value and riskto-benefit ratios associated with selective serotonin reuptake inhibitors (SSRIs), including concerns about quality of life in the long term (Almohammed et al., 2022) and conditions like Post-SSRI Sexual Dysfunction (Ben-Sheetrit et al., 2023), withdrawal syndrome (Horowitz and Taylor, 2019) and movement disorders (Revet et al., 2020).

The clinical reality is that, in many healthcare systems, demand outstrips the supply of evidence-based mental health treatment, which the recent COVID-19 pandemic has exacerbated (Zima et al., 2022). Therefore, the question of whether (and if so, which) easily accessible pharmaco-therapy options for the treatment of BSDs have comparable treatment effects to CBT interventions is clinically very relevant.

Several studies have examined the effects of pharmacological treatments for BN and BED. Antidepressants and other serotonin antagonists are the most widely studied drugs for the treatment of BN (Argyrou et al., 2023), while antidepressants, antiepileptics, psychostimulants, and antiobesity drugs have been studied for the treatment of BED (Monteleone et al., 2022).

A recent systematic meta-review (Monteleone et al., 2022) summarizes the comparisons and outcomes examined in a number of meta-analyses. Some have examined the efficacy of all forms of psychotherapy and pharmacotherapy (Bacaltchuk et al., 2001; Devlin, 1996; Fairburn and Hay, 1992; Ghaderi et al., 2018; Hilbert et al., 2019; Mcelroy et al., 2015; Mitchell et al., 1993; Peat et al., 2017; Ramacciotti et al., 2013; Slade et al., 2018; Svaldi et al., 2019; Vocks et al., 2010), but no study has directly compared the efficacy of CBT-based interventions with that of individual types of pharmacotherapy. Furthermore, it is often the case that a number of smaller and/or crossover pharmacotherapy trials are excluded from meta-analyses, especially for BN, due to the strict definition of remission that has often been utilized, i.e. 100% abstinence from symptoms for at least 2 weeks or no longer meeting the diagnostic criteria (Bacaltchuk and Hay, 2003; Williams et al., 2012). Finally, important clinical outcomes, including symptoms of anxiety and quality of life, have rarely been examined.

The present meta-analysis of randomized controlled trials (RCTs) aims to compare the effects of CBT interventions with any pharmacological treatment of BSDs for a wide range of important clinical outcomes. The goal is to address the clinical question of whether, and if so which, medication options could potentially be used as an alternative to CBT in treating these patient populations.

Materials and methods

An a priori written study protocol was published in PROSPERO (number: CRD42021230473) and is presented in Supplementary Material A. The systematic review and meta-analysis were conducted according to PRISMA (Preferred Reporting Items for Systematic reviews and Meta-Analyses) guidelines (Moher et al., 2009) (see Supplementary Material B: PRISMA checklist).

Participants and interventions

We included all RCTs that compared the effect of CBT (regardless of model or delivery format, including self-help CBT-based interventions) to any pharmacological treatment in patients with BN and/or BED, without restrictions regarding age, gender or comorbidities.

Search strategy and selection criteria

A comprehensive, systematic literature search was conducted in Embase, PubMed, Cochrane Central Register of Controlled Trials (CENTRAL) and ClinicalTrials.gov up to 15 February 2022 (last update 15 February 2023). The search strategy combined the following terms: ((OR (bulimi*) OR (binge) OR (binge*) OR (eating disorder) OR (eating disorder*)) AND ((OR (random*) OR (blind) OR (blind*) OR (clinical trial)). We also manually reviewed the reference lists of all included studies and reference lists of previously published reviews and meta-analyses.

There were no limitations in terms of language, date or publication status in the study selection process. At least two review authors worked independently to screen all the studies identified following de-duplication and assess the risk of bias using the Cochrane Risk of Bias Tool for Randomized Trials (Higgins et al., 2011). Conflicts at any stage were resolved through discussion with a third author. The first and/or corresponding authors of the original studies were contacted via email to obtain any missing data.

Outcome measures and data extraction

The primary outcomes were (1) remission, defined as a 100% reduction in bulimia/binge eating-related symptoms over a minimum of a 2-week period or as no longer meeting the relevant diagnostic criteria (including cognitive

elements) (Bacaltchuk and Hay, 2003; Williams et al., 2012), and (2) frequency of binges.

Additional outcomes were (1) frequency of purges (only in bulimia trials); (2) response, defined as a reduction of at least 50% in bulimic/binge-eating episodes or according to authors' definitions (Williams et al., 2012); (3) eating disorder psychopathology measured using the Eating Disorder Inventory (EDI) or any other validated scale whose psychometric properties have been documented in a peer-reviewed journal; (4) depressive symptoms, measured by validated rating scales such as the Hamilton Depression Rating Scale (HDRS) and Beck Depression Inventory (BDI); (5) anxiety symptoms, measured by validated rating scales such as the Hamilton Anxiety Rating Scale (HAM-A); (6) weight; (7) body mass index (BMI); (8) quality of life (QoL) measured by validated rating scales such as World Health Organization Ouality of Life Brief Version (WHOOOL-BREF). Ouality of Life Scale (QOLS), McGill Quality of Life Questionnaire (MQOL) or health-related quality of life (HRQOL); (9) total dropouts; (10) dropouts due to side effects; and (11) the total number of patients experiencing treatment-emergent adverse effects.

When authors of the original studies used imputation methods to account for missing data, these were preferred over completers' data. In the case of crossover trials, we used the first crossover phase, if possible, to avoid the problem of carryover effects. At least two review authors independently extracted data from all included studies, and any conflicts were resolved through discussion with a third author. Missing standard deviations were estimated from *p*-values or other information, or they were substituted by the mean standard deviation of the other included studies.

Statistical analysis

We performed a pairwise meta-analysis using the Review Manager 5.4 software (The Cochrane Collaboration, 2020). We employed a random-effects model (DerSimonian and Laird, 1986), as it is usually more conservative in terms of statistical significance, although, as a disadvantage, it puts added weight onto smaller studies, which can either inflate or deflate the effect size. We therefore examined whether using a fixed-effects model would markedly affect the results in a sensitivity analysis for the primary outcomes only.

All analyses were on an intention-to-treat (ITT) basis, whenever possible. Our preferred effect size for dichotomous outcomes was the risk ratio (RR). The effect size for continuous outcomes was the weighted mean difference (MD). If data were presented in different scales, then the standardized mean difference (SMD) was calculated instead. Missing standard deviations were calculated from standard errors or estimated from confidence intervals (CIs), *t*-values or *p*-values. Heterogeneity was investigated by inspection of the forest plots. Statistical heterogeneity was tested with the chisquare test and was quantified by calculating the I^2 statistics and its 95% CI. Potential reasons for heterogeneity were explored with subgroup analyses, but these were only conducted for the primary outcomes.

We planned several subgroup analyses considering the following a priori variables: individuals with BN vs individuals with BED; adults vs adolescents; previously treated participants vs not; participants with comorbidities vs no comorbidities; sponsored trials vs not; study duration up to 8 weeks vs longer; and treatment groups according to their main therapeutic concept. At the time the protocol was designed, the following sensitivity analyses were planned for the primary outcomes: exclusion of non-double-blind studies; exclusion of studies that did not use operationalized criteria to diagnose BN/BED; exclusion of studies that presented only completer analyses; exclusion of studies with high risk of bias; fixed effects instead of randomeffects model; and exclusion of studies with imputed data.

Results

Search

We identified 11 RCTs, which met our inclusion criteria. The studies were published from 1990 to 2019 and included 531 participants in total.

Description of included studies

The PRISMA flowchart is shown in Figure S1. Table 1 presents the details of all included studies. Notably, out of the 11 included studies, 10 took place in the United States (Agras et al., 1992; Devlin et al., 2005; Grilo et al., 2005, 2014; Leitenberg et al., 1994; Quilty et al., 2019; Walsh et al., 1997) and just one in Europe (Jacobi et al., 2002) (Hamburg, Germany). A total of 319 participants (60.07% of the individuals included in the analysis) were diagnosed with BN (7 studies) (Agras et al., 1992; Goldbloom et al., 1997; Jacobi et al., 2002; Leitenberg et al., 1994; Mitchell et al., 1990; Walsh et al., 1997, 2004). A total of 212 participants (39.93%) were diagnosed with BED (Devlin et al., 2005; Grilo et al., 2005, 2014; Quilty et al., 2019) (4 studies). The participants' mean age was 31.19 years. Two RCTs (Agras et al., 1992; Leitenberg et al., 1994) compared CBT (18–22 sessions) with desipramine (100–187.5 mg/day) for 20–24 weeks among patients diagnosed with BN (N=61). Five RCTs (Devlin et al., 2005; Goldbloom et al., 1997; Grilo et al., 2005; Jacobi et al., 2002; Walsh et al., 2004) compared CBT (6-20 sessions) with fluoxetine (40-60 mg/ day) for 16-20 weeks among patients with BN or BED (N=239). One RCT (Grilo et al., 2014) compared a selfhelp CBT intervention with sibutramine (15 mg/day) for 16 weeks among patients diagnosed with BED (N=51).

One RCT (Quilty et al., 2019) compared CBT (12 sessions) with methylphenidate (45 mg/day) for 12 weeks among patients with BED (N=49). In one RCT (Walsh et al., 1997), CBT (20 sessions) was compared with desipramine (188 mg/day) or fluoxetine (55 mg/day) therapy for 16 weeks. Finally, one RCT (Mitchell et al., 1990) compared CBT group therapy (12 weeks) with imipramine hydrochloride (300 mg/day) for 12 weeks among patients diagnosed with BN (N=78).

Risk of bias assessment

The results of the assessment of the risk of bias summary are presented in Figure 1 and the risk of bias graph in Figure S2. Strikingly, there was a high risk of bias in performance blinding among all included studies.

Remission and frequency of binge episodes (primary outcomes)

Figure 2 presents the results for the primary outcome remission. CBT was found to be superior to fluoxetine (RR=3.24, 95% CI=[1.46, 7.18], p=0.004, 2 RCTs, N=81, $I^2=0\%$). Comparisons of CBT with other drugs did not reveal any important differences, but only one RCT was available per comparison and sample sizes were very small, ranging from 9 to 53 patients. The pooled result for all antidepressants vs CBT favored CBT (RR=2.24, 95% CI=[1.03, 4.87], p=0.04, 4 RCTs, N=143, $I^2=29\%$, Figure S3).

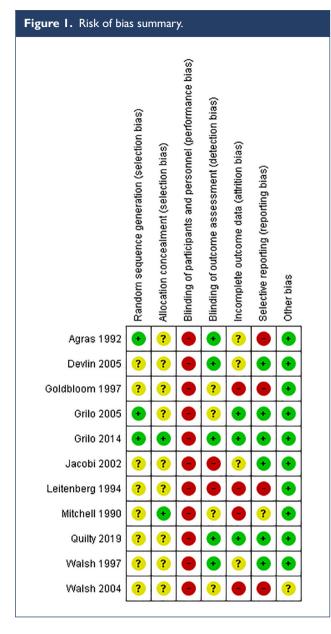
As for the frequency of binge episodes (Figure S4), CBT was found to be superior to imipramine (SMD=-0.93, 95% CI=[-1.40, -0.46], p < 0.01, 1 RCT, N=78), and there was a trend for superiority of CBT over desipramine (SMD=-0.49, 95% CI=[-1.07, 0.09], p=0.10, 1 RCT, N=47), and methylphenidate (SMD=-0.56, 95% CI=[-1.13, 0.02], p=0.06, 1 RCT, N=49). The pooled result for all antidepressants vs CBT favored CBT marginally (SMD=-0.35, 95% CI=[-0.69, -0.01], p=0.04, 8 RCTs, N=396, $l^2=63\%$, Figure S5).

CBT was found to be more efficacious than imipramine for the outcomes frequency of purge (Figure 3; SMD=-0.93, 95% CI=[-1.40, -0.45], p < 0.01, 1 RCT, N=78), eating disorder psychopathology (Figure S6; SMD=-1.02, 95% CI=[-1.54, -0.50], p < 0.01, 1 RCT, N=65), and dropouts due to any reason (Figure S7; RR=0.35, 95% CI=[0.15, 0.82], p=0.02, 1 RCT, N=88). When examining all antidepressants as a subgroup, CBT was found to be more efficacious only for the outcome eating disorder psychopathology (Figure S8; SMD=-0.64, 95% CI=[-0.99, -0.29], p < 0.01, 5 RCTs, N=239).

With regard to the outcome BMI, CBT had higher BMI values compared to methlylphenidate (Figure S9; MD=5.78, 95% CI=[1.37, 10.19], p=0.01, 1 RCT, N=49) and lower compared to fluoxetine (Figure S9; MD=-2.71, 95% CI=[-5.33, -0.10], p=0.04, 2 RCTs, N=100).

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Buttorinde buttorinde240.8-33N=34N=34Support Buttorinde ButtorindeSupport Buttorinde ButtorindeSupport Buttorinde ButtorindeSupport Buttorinde ButtorindeSupport Buttorinde ButtorindeSupport Buttorinde Buttori	Study ID	Study design	Treatment duration (weeks)	Mean age of participants (years)	No. randomized on CBT and number of sessions	No. randomized on pharmacotherapy. drug and mean dose per day	Main outcomes	Definition of response/remission to treatment	Eating disorder diagnostic criteria
Mediation building1013 \mathbb{V}^{-13} $$	Agras et al. (1992)	Randomized, parallel, single- blind	24	29.6	N = 23 18 sessions	N = 24 Desipramine 187.5 mg	Frequency of binge eating and purge episodes	Remission from binge eating and purge behaviors without specifying for how long	Bulimia nervosa (DSM-III-R)
Wardonized parallel, open burdenized,1N=14 burdenized burdenized burdenized,N=14 burdenized burdenized,Entry Diorden Examination (ED)Wardonized, burdenized, 	Devlin (2005)	Randomized, parallel, double- blind ^a	20	43.0	N = 25 20 sessions	N = 32 Fluoxetine 40mg	Frequency of binge eating episodes Weight	Remission from binge eating the previous 4 weeks	Binge eating disorder (DSM-IV)
Kardenized, bined101-30N=3N=27Frequency of bing exting binoreMardenized, bined10 <t< td=""><td>Goldbloom et al. (1997)</td><td>Randomized, parallel, open</td><td>9</td><td>25.8</td><td>N = 14 16 sessions</td><td>N = 12 Fluoxetine 60 mg</td><td>Eating Disorder Examination (EDE)</td><td>Remission from binge eating and vomiting 4 weeks post trattment Response: 50% reduction of binge eating and vomiting 4 weeks post treatment</td><td>Bulimia nervosa (DSM-III-R)</td></t<>	Goldbloom et al. (1997)	Randomized, parallel, open	9	25.8	N = 14 16 sessions	N = 12 Fluoxetine 60 mg	Eating Disorder Examination (EDE)	Remission from binge eating and vomiting 4 weeks post trattment Response: 50% reduction of binge eating and vomiting 4 weeks post treatment	Bulimia nervosa (DSM-III-R)
Radomized paralel164.34N=35 celiniap CeT book pisotianN=26 celiniap cet of objects biotramine biotramineN=26 celiniap cet of objects biotramine biotramineN=26 celiniap cet of objects biotramine celiniap cet of objectsTequeroy ob large action dispication weight celiniap celiniap cet of objectsRadomized paralel, open1620N=19 celiniap celiniapN=16 celiniap celiniap celiniapN=16 celiniap celiniap celiniapN=16 celiniap celiniap celiniapN=16 celiniap celiniap celiniap celiniapN=16 celiniap celiniap celiniap celiniapN=16 celiniap celiniap celiniap celiniapN=16 celiniap celiniap celiniap celiniap celiniapN=16 celiniap celiniap celiniap celiniap celiniap celiniap celiniap celiniapN=16 celiniap celiniap 	Grilo et al. (2005)	Randomized, parallel, double- blind ^a	9	43.9	N = 28 16 sessions	N = 27 Fluoxetine 60 mg	Frequency of binge eating episodes	No binge eating episode the last 28 days	Binge eating disorder (DSM-IV)
Randomized parallel open parallel open160.8-16N=16Fleuxetine fleuxetine domg domg parallel openFleuxetine fleuxetine fleuxetine domgFleuxetine fleuxetine fleuxetine domgFleuxetine fleuxetine fleuxetine fleuxetine fleuxetineFleuxetine fleuxetine fleuxetine fleuxetineFleuxetine fleuxetine fleuxetine fleuxetineFleuxetine fleuxetine fleuxetine fleuxetine fleuxetineFleuxetine fleuxetine 	Grilo et al. (2014)	Randomized, parallel, double- blind ^a	9	43.4		N=26 Sibutramine 15 mg	Frequency of binge eating episodes Weight loss	No binge eating episode the last month based on EDE-I	Binge eating disorder (DSM-5)
Randomixed, parallel, open2036.7N=7N=7N=7Randomixed, parallel, double12.3 sessionsDesipramine toomsIncervy of binge exting toomsRandomixed, blind*12.3 sessionsN=45Desipramine toomsIncervy of binge exting toppe extingRandomixed, blind*122.3 sessionsN=37N=45Interview of binge exting toppe extingRandomixed, blind*122.3 sessionsN=27N=27Interview of binge exting toppe extingRandomixed, blind*102.3 sessionsN=27N=22Interview of binge exting toppe extingRandomixed, blind*102.3 sessionsN=28N=27Interview of binge exting toppe extingRandomixed, blind*162.5 0N=28N=28N=28Randomixed, blind*162.5 0SossionsDesipramine and then Fluxoetine (for toppe extingRandomixed, blind*162.5 0N=28N=28Randomixed, blind*162.5 0SossionsSossionsRandomixed, blind*162.5 0N=28N=28Randomixed, blind*162.5 0N=28N=28Randomixed, blind*162.5 0SossionsSossionsRandomixed, blind*162.5 0N=28N=28Randomixed, blind*162.5 0N=28N=28Randomixed, blind*162.5 0SossionsSocsionsRandomixed	Jacobi et al. (2002)	Randomized, parallel, open	9	26.0	N = 19 20 sessions	N = 16 Fluoxetine 40 mg	Frequency of binge eating and purge episodes EDI and Three Factor Eating Questionnaire (TFEQ) score	Remission from binge eating and purge behaviors without specifying for how long	Bulimia nervosa (DSM-III-R)
Randomized, parallel, double- blind*1223.5N=33 htensive group CBT hydrochloride 300mgN=45 hydrochloride hydrochlorideFrequery of binge exting and purge episodes (self- hydrochloride minamine menamineRandomized, parallel, open123.2.8N=27 (2 sesions)N=22 (2 sesions)N=22 (2 sesions)N=22 (2 sesions)N=22 (2 sesions)N=23 (2 sesions)N=23 	Leitenberg et al. (1994)	Randomized, parallel, open	20	26.7	N=7 22 sessions	N=7 Desipramine 100 mg	Frequency of binge eating and purge episodes	Remission from vomiting without specifying for how long	Bulimia nervosa (DSM-III-R)
Randomized, parallel, open1232.8N=27 tequency of binge eating 45 mgN=22 methylphenidate 45 mgN=23 methylphenidate 45 mgN=23 tequency of binge eating episodesRandomized, parallel, double blind*1625.0N=25 20 sessionsN=28 Designamine and then Fluoxetine (for parallel single blind*Frequency of binge eating episodesRandomized, blind*1625.0N=25 20 sessionsN=28 Designamine and then Fluoxetine (for parallel single blind*Frequency of binge eating and vomiting serviceRandomized, blind*1630.6N=25 BB/B/B designamine/S5/B fluoxetine 	Mitchell et al. (1990)	Randomized, parallel, double- blind ^a	12	23.5	N = 33 Intensive group CBT 12 weeks	N = 45 Imipramine Hydrochloride 300mg	Frequency of binge eating and purge episodes (self- induced vomiting)	Remission from bulimic episodes: no episodes for the last 2 weeks	Bulimia nervosa (DSM-III)
Randomized, parallel, double- blind ⁿ 16 25.0 N=25 N=28 Designamine and then Fluoxetine (for and vomiting participants that designamine was not effective) Frequency of binge eating and vomiting participants that designamine/55 mg fluoxetine Randomized, blind 16 30.6 -8 N=20 Randomized, blind 16 30.6 -8 Fluoxetine Randomized, blind 16 30.6 -8 N=20 Randomized, blind 16 30.6 -8 Fluoxetine Sessions (2 optional), guided self-help CBT 60mg and vomiting	Quilty et al. (2019)	Randomized, parallel, open	12	32.8	N = 27 12 sessions	N = 22 Methylphenidate 45 mg	Frequency of binge eating episodes	No binge eating episode for the previous 4 weeks of treatment	Binge eating disorder (DSM-5)
Randomized, 16 30.6 N=25 N=20 Frequency of binge eating eating parallel, single- 6-8 Fluoxetine and vomiting blind sessions (2 optional), 60 mg guided self-help CBT guided self-help CBT	Walsh et al. (1997)	Randomized, parallel, double- blind ^a	9	25.0	N=25 20 sessions	N=28 Desipramine and then Fluoxetine (for participants that desipramine was not effective) 188mg desipramine/55mg fluoxetine	Frequency of binge eating and vomiting	Remission from binge eating episodes and/or purge behaviors the two last weeks	Bulimia nervosa (DSM-III-R)
	Walsh et al. (2004)	Randomized, parallel, single- blind	16	30.6	N=25 6-8 sessions (2 optional), guided self-help CBT	N = 20 Fluoxetine 60 mg	Frequency of binge eating and vomiting	Remission from binge eating/purging episodes for the previous 4 weeks of treatment	Bulimia nervosa (DSM-IV)

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There were no statistically significant differences for any of the individual CBT vs drug comparisons in terms of response (Figure S10; RR=1.47, 95% CI=[0.98, 2.21], p=0.06, 5 RCTs, N=170), depression (Figure S11; MD=-0.70, 95% CI=[-2.38, 0.98], p=0.41, 8 RCTs, N=388), anxiety (Figure S12; MD=-0.07, 95% CI=[-0.62, 0.48], p=0.81, 2 RCTs, N=131), weight (Figure S13; SMD=0.15, 95% CI=[-0.30, 0.60], p=0.50, 4 RCTs, N=179), quality of life (Figure S14; MD=0.57, 95% CI=[-0.16, 1.30], p=0.13, 1 RCT, N=49), and dropouts due to side effects (Figure S15; RR=0.37, 95% CI=[0.03, 5.38], p=0.47, 2 RCTs, N=113).

Publication bias

Data from fewer than 10 studies were available for most of the outcomes. When fewer than 10 studies are included in a

meta-analysis, asymmetry testing for funnel plots should not be employed (Egger et al., 1997; Higgins et al., 2019). As a result, we were unable to use funnel plots for the outcomes of remission, frequency of purges, response, eating disorder psychopathology, depressive and anxiety symptoms, weight, BMI, quality of life, total dropouts and dropouts due to side effects. However, the meta-analysis included 10 studies examining the outcome of frequency of binges, and in this case, the funnel plot was symmetrical (Figure 4).

Discussion

Summary and contributions

CBT is considered the first-line treatment option for BSDs (Treasure et al., 2020). Nevertheless, this treatment modality is underutilized and may be difficult to access in everyday clinical practice. Exploring other therapeutic alternatives is important. Thus, we attempted to answer the important clinical question as to whether (and if so, which) easily accessible pharmacotherapy options could potentially be equally effective to CBT for the treatment of BSDs.

Eleven RCTs published from 1990 to 2019 were identified, with a total of 531 participants. They compared CBT with five drugs, namely fluoxetine, imipramine, desipramine, methylphenidate and sibutramine. Our findings do not allow us to recommend specific drugs as an alternative to CBT, as few statistically significant differences were identified in the examined outcomes.

Remission, defined as a 100% reduction in bulimia/ binge eating-related symptoms over a minimum of a 2-week period or as no longer meeting the relevant diagnostic criteria (Bacaltchuk and Hay, 2003; Williams et al., 2012), was examined in 6 out of the 11 included RCTs. CBT was superior to fluoxetine (dose range, 40–60 mg) and the pooled data for antidepressants, but few studies were included, and the sample sizes were small.

As full remission is often not achieved or maintained in the long term (Smink et al., 2013; Treasure et al., 2020), response to treatment (also referred to as partial remission) is an outcome that perhaps resonates better with the clinician and the service-user. We expected that different trials would use different definitions of response. However, it has been shown that, as long as relative measures of risk are applied, meta-analytic results do not differ significantly depending on the exact cut-off applied (Furukawa et al., 2011). We prioritized using the definition of response as a reduction of at least 50% in bulimic/binge-eating episodes (Williams et al., 2012), where data were available. When a variety of definitions were provided and this criterion was not met, we used the authors' definitions of response, selecting the strictest. Although no significant difference between CBT and antidepressants was shown, there was a trend in favor of CBT (response rates of 44% vs 29%).

Figure 2. Remission.

	CBT		Medica	lion		Risk Ratio	Risk Ratio
Study or Subgroup					Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.1.1 CBT vs Desipram		Total	Lvents	Total	weight	m-n, Kanuom, 55% Ci	M-H, Randolli, 95% Cl
Leitenberg 1994	5	6	0	3	4.5%	6.29 [0.46, 86.47]	
Subtotal (95% CI)	5	6	0	3	4.5%	6.29 [0.46, 86.47]	
Total events	5	·	0		non	0.20 [0.10, 00111]	
Heterogeneity: Not appl	-		0				
Test for overall effect: Z		P = 0.1	7)				
			.,				
1.1.2 CBT vs Desiprami	ine/Fluo	xetine					
Walsh 1997	5	25	6	28	16.8%	0.93 [0.32, 2.69]	_
Subtotal (95% CI)		25		28	16.8%	0.93 [0.32, 2.69]	
Total events	5		6				
Heterogeneity: Not appl							
Test for overall effect: Z	= 0.13 (P = 0.9	0)				
1.1.3 CBT vs Fluoxetine							
Goldbloom 1997	6	14	2	12	11.9%	2.57 [0.63, 10.45]	
Grilo 2005	15	28	4	27	18.4%	3.62 [1.37, 9.52]	
Subtotal (95% CI)		42		39	30.4%	3.24 [1.46, 7.18]	
Total events	21		6				
Heterogeneity: Tau ² = 0	.00; Chi	² = 0.1	5, df = 1 (F	P = 0.69	3); I² = 0%		
Test for overall effect: Z	= 2.89 (P = 0.0	04)				
1.1.4 CBT vs Methylphe	nidate						
Quilty 2019	16	27	10	22	27.6%	1.30 [0.75, 2.27]	
Subtotal (95% CI)	10	27	10	22	27.6%	1.30 [0.75, 2.27]	•
Total events	16		10				-
Heterogeneity: Not appl	icable						
Test for overall effect: Z		P = 0.3	5)				
1.1.5 CBT vs Sibutrami	ne						
Grilo 2014	6	25	10	26	20.8%	0.62 [0.27, 1.46]	
Subtotal (95% CI)		25		26	20.8%	0.62 [0.27, 1.46]	
Total events	6		10				
Heterogeneity: Not appl			-				
Test for overall effect: Z	= 1.09 (P = 0.2	:8)				
Total (95% CI)		125		118	100.0%	1.48 [0.82, 2.68]	◆
Total events	53		32				
Heterogeneity: Tau ² = 0	.25; Chi	² = 10.0	00, df = 5	(P = 0.0	08); I ² = 51	0%	0.01 0.1 1 10 100
Test for overall effect: Z				1211	10.00		0.01 0.1 1 10 100 Favours Medication Favours CBT
Test for subaroup differ	ences. I	$Chi^2 = 9$	= 1h 08 F	4 (P = 1)	105) F=	58 3%	

However, it is worth noting that only desipramine and fluoxetine were examined among antidepressants, and the availability of data was very limited.

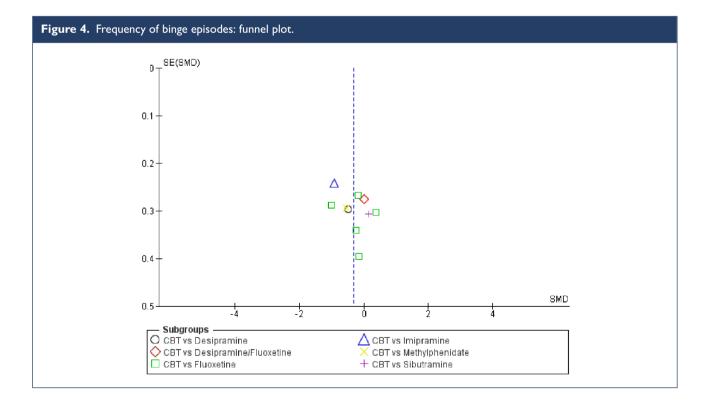
Reduction of frequency of binges is another commonly used outcome measure. Our findings indicated a superiority for CBT over imipramine and over antidepressants in general, albeit with a relatively small effect size (SMD=-0.35, 95% CI=[-0.69, -0.01], 8 RCTs, N=396). This finding is in line with previously published work comparing all forms of psychotherapy to pharmacotherapy (Svaldi et al., 2019; Vocks et al., 2010). The clinical significance of the frequency of binges, however, remains controversial as it does not seem to correlate with long-term treatment outcomes such as weight loss (Delinsky et al., 2006), and over-valuation of weight/shape and body dissatisfaction are perhaps better clinical indicators of the severity of illness (Grilo et al., 2015; Wilson and Sysko, 2009) and predictors of suicidality (Perkins and Brausch, 2019; Rufino et al., 2018).

Purging behavior is associated with significant medical comorbidities and healthcare utilization in the short term and long term (Mehler et al., 2015). We found that CBT was superior to imipramine and tended to be superior to desipramine. However, no significant differences were found between CBT and fluoxetine or CBT vs antidepressants overall.

Severity of psychopathology is an important indicator of illness severity and functional decline (Dahlgren et al., 2017). Therefore, this outcome is clinically relevant. We found that CBT was superior to imipramine and antidepressants overall for improving psychopathology, as measured by the Eating Disorder Examination Global Score, with a medium to large effect size (SMD=-0.64, 95% CI=[-0.99,

Figure 3. Frequency of purges.

Study or Subgroup Mean SD Total Mean SD Total Meight N, Random, 95% CI 1.3.1 CBT vs Designamine Agras 1992 1.7 2.7 23 4.3 6.7 24 15.5% -0.50 [-1.08, 0.08] Leitenberg 1994 1.43 3.3 7 7.14 4.24 7 8.7% -1.41 [-2.62, -0.19] Subtotal (95% CI) 30 31 24.3% -0.79 [-1.62, 0.04] Image: the state			CBT		Medication			Std. Mean Difference		Std. Mean Difference
Agras 1992 1.7 2.7 23 4.3 6.7 24 15.5% -0.50 [-1.08, 0.08] Leitenberg 1994 1.43 3.3 7 7.14 4.24 7 8.7% -1.41 [-2.62, -0.19] Subtotal (95% CI) 30 31 24.3% -0.79 [-1.62, 0.04] Heterogeneity: Tau ² = 0.18; Chi ² = 1.76, df = 1 ($P = 0.19$); $P = 43\%$ Test for overall effect: $Z = 1.86$ ($P = 0.06$) 1.3.2 CBT vs Designamine/Fluoxetine Walsh 1997 5.6 15 25 3.7 5 28 16.0% 0.17 [-0.37, 0.71] Butotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect: $Z = 0.62$ ($P = 0.53$) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 ($P = 0.18$); $P = 42\%$ Test for overall effect: $Z = 0.65$ ($P = 0.52$) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% CI) 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ ($P = 0.0001$) Total (95% CI) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, df = 6 ($P = 0.0008$); $P = 74\%$	Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Leitenberg 1994 1.43 3.3 7 7.14 4.24 7 8.7% $-1.41[2.62, -0.19]$ Subtotal (95% CI) 30 31 24.3% $-0.79[-1.62, 0.04]$ Heterogeneity: Tau ² = 0.18; Ch ² = 1.76, df = 1 (P = 0.19); l ² = 43% Test for overall effect Z = 1.86 (P = 0.06) 1.3.2 CBT vs Designamine/Fluoxetine Walsh 1997 5.6 15 25 3.7 5 28 16.0% 0.17 [-0.37, 0.71] Subtotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect Z = 0.62 (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% $-0.36 [-1.14, 0.42]$ Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); l ² = 42% Test for overall effect Z = 0.65 (P = 0.52) 1.3.4 CBT vs Impramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% $-0.93 [-1.40, -0.45]$ Subtotal (95% CI) 33 45 16.8% $-0.93 [-1.40, -0.45]$ Heterogeneity: Not applicable Test for overall effect Z = 3.84 (P = 0.0001) Total (95% CI) 146 152 100.0% $-0.27 [-0.74, 0.21]$	1.3.1 CBT vs Desipra	mine								
Subtotal (95% CI) 30 31 24.3% -0.79 [-1.62, 0.04] Heterogeneity: Tau ² = 0.18; Chi ² = 1.76, df = 1 (P = 0.19); I ² = 43% Test for overall effect: $Z = 1.86$ (P = 0.06) 1.3.2 CBT vs Designamine/Fluoxetine Walsh 1997 5.6 15 25 3.7 5 28 16.0% 0.17 [-0.37, 0.71] Subtotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect: $Z = 0.62$ (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 18.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); I ² = 42% Test for overall effect: $Z = 0.85$ (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ (P = 0.0001) Total (95% CI) 16 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, df = 6 (P = 0.0008); I ² = 74% -4 -2 0 2 4	Agras 1992	1.7	2.7	23	4.3	6.7	24	15.5%	-0.50 [-1.08, 0.08]	
Heterogeneity: Tau ² = 0.18; Chi ² = 1.76, df = 1 (P = 0.19); ² = 43% Test for overall effect: Z = 1.86 (P = 0.06) 1.3.2 CBT vs Desigramine . Fluoxetine Walsh 1997 5.6 15 25 3.7 5 28 16.0% 0.17 [-0.37, 0.71] Subtotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect: Z = 0.62 (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); ² = 42% Test for overall effect: Z = 0.85 (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% CI) 33 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: Z = 3.84 (P = 0.0001) Total (95% CI) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, dif = 6 (P = 0.0008); ² = 74%	Leitenberg 1994	1.43	3.3		7.14	4.24	7	8.7%	-1.41 [-2.62, -0.19]	
Test for overall effect: $Z = 1.86$ (P = 0.06) 1.3.2 CBT vs Designamine [Fluoxetine Walsh 1997 5.6 15 25 3.7 5 28 16.0% 0.17 [-0.37, 0.71] Subtotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect: $Z = 0.62$ (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); ² = 42% Test for overall effect: $Z = 0.65$ (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% CI) 33 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ (P = 0.0001) Total (95% CI) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, dif = 6 (P = 0.0008); ² = 74%	Subtotal (95% CI)			30			31	24.3%	-0.79 [-1.62, 0.04]	
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Subtotal (95% CI) 25 28 16.0% 0.17 [-0.37, 0.71] Heterogeneity: Not applicable Test for overall effect: $Z = 0.62$ (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); l ² = 42% Test for overall effect: $Z = 0.65$ (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% CI) 33 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ (P = 0.0001) Total (95% CI) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, df = 6 (P = 0.0008); l ² = 74%	1.3.2 CBT vs Desipra	mine/Flu	oxetine	•						
Heterogeneity: Not applicable Test for overall effect: $Z = 0.62$ (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% Cl) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); i ² = 42% Test for overall effect: $Z = 0.65$ (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% Cl) 33 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ (P = 0.0001) Total (95% Cl) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, df = 6 (P = 0.0008); i ² = 74%					3.7	5	28	16.0%	0.17 [-0.37, 0.71]	
Test for overall effect: $Z = 0.62$ (P = 0.53) 1.3.3 CBT vs Fluoxetine Goldbloom 1997 9 16.8 14 17.3 27.2 12 13.1% -0.36 [-1.14, 0.42] Jacobi 2002 25.4 35 19 20.1 27.1 16 14.5% 0.16 [-0.50, 0.83] Walsh 2004 46.12 56.75 25 19.85 25.8 20 15.3% 0.56 [-0.04, 1.16] Subtotal (95% CI) 58 48 42.9% 0.17 [-0.34, 0.68] Heterogeneity: Tau ² = 0.09; Chi ² = 3.43, df = 2 (P = 0.18); l ² = 42% Test for overall effect: $Z = 0.65$ (P = 0.52) 1.3.4 CBT vs Imipramine Mitchell 1990 -11.2 7.8 33 -3.9 7.8 45 16.8% -0.93 [-1.40, -0.45] Subtotal (95% CI) 33 45 16.8% -0.93 [-1.40, -0.45] Heterogeneity: Not applicable Test for overall effect: $Z = 3.84$ (P = 0.0001) Total (95% CI) 146 152 100.0% -0.27 [-0.74, 0.21] Heterogeneity: Tau ² = 0.29; Chi ² = 22.99, df = 6 (P = 0.0008); l ² = 74%	Subtotal (95% CI)			25			28	16.0%	0.17 [-0.37, 0.71]	
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-0.29], 5 RCTs, N=239). This finding supports the recommendations offered by major guidelines (Crone et al., 2023; NICE Guideline, 2017). Notably, other less popular treatment options, including methylphenidate and sibutramine, were found to be equally efficacious to CBT, but usable data were only available from one RCT per comparison. Further research into these treatment options is therefore needed.

Weight loss and BMI represent clinically important outcomes consistently reported in most RCTs. BSDs often cooccur with obesity (30-45%) and its associated metabolic consequences, well-known predictors of high cardiovascular risk and overall mortality (da Luz et al., 2018; Hay et al., 2015; Kessler et al., 2013). Furthermore, body weight, body shape and attitudes toward these factors constitute major precipitating and perpetuating elements in BSDs (Treasure et al., 2020). Obesity or being overweight plays a central role in perpetuating body uneasiness (Rotella et al., 2013), although other factors, such as relentless body shaming and the prevalence of a patriarchal diet culture within Western society, can also significantly contribute to body uneasiness. Individuals with higher body weight affected by eating disorders often fall victim to systemic discrimination related to their body shape, including weight stigma (Lawrence et al., 2022). Regrettably, weight stigma is prevalent even within the medical profession, resulting in individuals with higher weight experiencing medical gaslighting and trauma, highlighting the urgent need for substantial changes in healthcare and beyond (Batterham, 2022; Talumaa et al., 2022). Collectively, these factors likely contribute to increased stress and psychological distress. Unfortunately, the evidence supporting the efficacy of CBT for weight management in overweight/obese individuals with BSDs is weak (Palavras et al., 2017), and our analyses, with only a limited number of RCTs, did not permit any meaningful interpretation of the results.

Overall, our research review holds significant importance in advancing the field of eating disorders for several compelling reasons. First, we provide a thorough and upto-date analysis of all available evidence and incorporate a greater number of studies, enhancing the depth and relevance of our research. Second, our review expands its scope to encompass patients with BED, a relatively recently recognized diagnosis, thereby making our findings more pertinent and applicable to a broader population of individuals seeking treatment. Significantly, our study challenges the conventional perception that CBT stands as the undisputed gold standard for treating eating disorders, particularly in light of limited supporting evidence. By advocating for a wider range of treatment options, we aim to stimulate a shift in clinical practice and the development of local guidelines, potentially reducing the underutilization of pharmacotherapy in suitable cases, while also inspiring further primary clinical research in the field.

Limitations

This study has several limitations. First, and perhaps more importantly, a very small number of trials and consequently participants were included in the present meta-analysis. This did not allow us to perform the a priori planned subgroup and sensitivity analyses. Moreover, single agent results (with the possible exception of fluoxetine) are difficult to interpret adequately since many analyses included only one study.

According to Trikalinos et al., effect sizes are markedly changeable when less than 1000 participants have been included in a psychopharmacotherapy meta-analysis, but they seem to stabilize after this threshold has been reached (Trikalinos et al., 2004). In our work, none of the comparisons reached this threshold, and all sample sizes were minimal (ranging from 14 to 78 participants). This renders the derived evidence uncertain and does not allow for a definitive clinical interpretation.

In addition, the absence of several newer antidepressants and novel agents, such as lisdexamfetamine (which received FDA approval for BED in 2015), is noteworthy. Since 2005, there has been no recent trial comparing CBT to pharmacotherapy. Conducting such trials, particularly involving these novel agents, has become imperative to effectively inform clinical practice.

Moreover, blinding in psychotherapeutic interventions is always problematic, leading to performance and detection biases (Juul et al., 2020), which are strikingly present in all studies included in our meta-analysis. In addition, a significant degree of variation was evident in the delivery methods and dosages of CBT. This spectrum encompassed interventions ranging from 6 to 22 sessions and from guided self-help to group CBT and individual therapy. Addressing the placebo effect also presents challenges; factors such as time spent with the patient, empathy, and the therapeutic relationship hold substantial influence over treatment outcomes (Enck and Zipfel, 2019). In contrast, five of the identified RCTs employed a placebo-controlled doubleblind design specifically for drug assignment and management. All these diverse variables increase the risk of biased treatment estimates and may, to some extent, account for the identified effects.

Another limitation is that we decided to group BN and BED together. There are distinct differences between the two disorders, highlighted across the major classification systems (Claudino et al., 2019; Smith et al., 2017). However, BED and BN do share the majority of their biological, psychological, psychosocial and behavioral etiological factors (Treasure et al., 2020), and there are compelling arguments for a spectrum model (Brooks et al., 2012; Treasure et al., 2020). Furthermore, there are many similarities in terms of psychopathology, physical findings and their effects on functioning and quality of life, and there is a significant overlap in terms of treatment goals and

approaches (Treasure et al., 2020), which we believe are reflected in the outcomes examined in this work.

Future work

Our findings only reflect short-term treatment effects and do not provide any information regarding medium- and long-term efficacy, tolerability and acceptability. Evaluating long-term efficacy necessitates extended follow-up periods, but practical challenges such as limited resources and participant attrition often hinder their implementation, explaining the scarcity of available data. This scarcity extends across various mental health conditions, where limited and sometimes contradictory evidence regarding the long-term effects of CBT exists (Cuijpers et al., 2023; Hofmann et al., 2012; Sharma et al., 2021; Van Dis et al., 2020). However, this is a significant consideration given that many mental health conditions typically follow a chronic course.

For individuals with BSDs, there is a limited body of evidence supporting the medium- and long-term efficacy of CBT in maintaining remission (Agras, 1997; Carter and Fairburn, 1998; Wilfley et al., 2002). Similarly, the maintenance of therapeutic benefits in pharmacotherapy trials has rarely been examined. A relevant published placebo-controlled trial (Stunkard et al., 1996) indicates a high risk of relapse in terms of binge episodes following the discontinuation of medication at the 4-month follow-up among patients with BED. The long-term treatment outcomes of CBT compared to pharmacotherapy for BSDs remain poorly understood, underscoring the pressing need for further research in this critical area, which represents a major clinical priority.

Conclusion

Few RCTs have investigated the efficacy and safety of CBT interventions compared to any form of pharmacotherapy for the treatment of BSDs. Only a select few antidepressants, such as fluoxetine, imipramine and desipramine, have been examined. Our results are derived from a small number of participants per individual comparison, often relying on just one RCT.

Perhaps the most important finding of this work is that, due to the underpowered comparisons and the inherent methodological limitations of psychotherapy trials, more head-to-head trials comparing CBT with monotherapy drug interventions are needed to establish the superiority of one option over the other. Further research into pharmacotherapy, especially exploring newer antidepressants and novel agents, is warranted. In addition, combination trials examining pharmacotherapy options as an augmentation to CBT are lacking and would be a valuable addition to eating disorders research and clinical practice.

Author Contributions

M.T.S. and M.C. contributed to study conception and design; N.M., A.A., E.M., D.R.B., E.T. and Z.A.P. contributed to data collection; M.T.S., N.M. and A.S.L. contributed to analysis and interpretation of results; M.T.S. and A.S.L. contributed to draft manuscript preparation; N.C., G.P. and M.C. contributed to writing review and editing. All authors reviewed the results and approved the final version of the manuscript. All authors have read and agreed to the published version of the manuscript.

Declaration of Conflicting Interests

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Supplemental Material

Supplemental material for this article is available online.

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