

## Titel: Quantification of model risk in quadratic hedging in finance

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ABSTRACT. In this paper the effect of the choice of the model on partial hedging in incomplete markets in finance is estimated. In fact we compare the quadratic hedging strategies in a martingale setting for a claim when two models for the underlying stock price are considered. The first model is a geometric Lévy process in which the small jumps might have infinite activity. The second model is a geometric Lévy process where the small jumps are replaced by a Brownian motion which is appropriately scaled. The hedging strategies are related to solutions of backward stochastic differential equations with jumps which are driven by a Brownian motion and a Poisson random measure to prove that they are robust towards the choice of the model for the market prices and to estimate the model risk.

### 1. Introduction

In financial markets, the hedging of derivatives is in general set in the non-arbitrage framework and can be technically performed under a related pricing measure that is a risk-neutral measure. Under this measure the discounted prices of the underlying primaries are martingales. In some markets, for example, in the context of energy derivatives, the underlying, electricity, cannot be stored. Hence hedging does not require that the pricing measure is a risk-neutral measure. See e.g., Benth et al. (2008) for more details. In this case the discounted stock price process is a semimartingale under the pricing measure.

When jumps are present in the model for the stock price, the market is in general incomplete and there is no self-financing hedging strategy which allows to attain the contingent claim at maturity. In other words, one cannot eliminate the risk completely. However it is possible to find 'partial' hedging strategies which minimise some risk. One way to determine these 'partial' hedging strategies is to introduce a subjective criterion according to which strategies are optimised.

We consider two types of quadratic hedging strategies. The first is called risk-minimisation (RM) in the martingale setting and local risk-minimisation (LRM) in the semimartingale setting. These strategies replicate the option's payoff, but they are not self-financing (see, e.g., Schweizer (2001)). In the martingale setting the RM strategies minimise the risk process which is induced by the fact that the strategy is not self-financing. In the semimartingale setting the LRM strategies minimise the risk in a 'local' sense (see Schweizer (1988, 1991)). In the second approach, called mean-variance hedging (MVH), the strategy is self-financing and one minimises the quadratic hedging error at maturity in mean square sense (see, e.g., Schweizer (2001)). These strategies are related to the study of the Galtchouk-Kunita-Watanabe (GKW) or/and the Föllmer-Schweizer decomposition (FS) decomposition which are both backward stochastic differential equations (BSDEJ) (see Choulli et al. (2010) for more about these decompositions). BSDEJs have important applications in mathematical finance. In particular, they play a major role in hedging and in nonlinear pricing theory for incomplete markets (see El Karoui et al. (1997) for an overview).

The aim in this paper is to investigate whether these quadratic hedging strategies (RM and MVH) in incomplete markets in finance are robust to the variation of the model in the



martingale setting. In our analysis we follow the approach by Di Nunno et al. (2015), who dealt with the semimartingale setting, to write the value of the discounted portfolio in a RM strategy as a solution to a backward stochastic differential equation with jumps (BSDEJ). The main difference to the present paper lies in the fact that we have to deal with the martingale measure which depends on the choice of the model for the asset price process.

To model the asset price dynamics we consider two geometric Lévy processes. This class of models describes well realistic asset price dynamics and is well established in the literature (see e.g., Cont and Tankov (2004)). The first model  $(S_t)_{t\in[0,T]}$  is driven by a Lévy process in which the small jumps might have infinite activity. The second model  $(S_t^{\varepsilon})_{t\in[0,T]}$  is driven by a Lévy process in which we replace the jumps with absolute size smaller than  $\varepsilon > 0$  by an appropriately scaled Brownian motion. This idea of shifting from a model with small jumps to another where these variations are represented by some appropriately scaled continuous component goes back to Asmussen and Rosinski (2001) who proved that the second model approximates the first one.

The process  $(S_t^{\varepsilon})_{t\in[0,T]}$  is useful from a simulation point of view. Indeed, it contains a compound Poisson process and a scaled Brownian motion which are both easy to simulate. Whereas it is not easy to simulate the infinite activity of the small jumps in the process  $(S_t)_{t\in[0,T]}$  (see Cont and Tankov (2004) for more about simulation of Lévy processes).

However, the interest of this paper is model risk. From a modelling point of view, we may think of two financial agents who want to price and hedge an option. One is considering  $(S_t)_{t\in[0,T]}$  as a model for the price process and the other is considering  $(S_t^{\varepsilon})_{t\in[0,T]}$ . In other words, the first agent chooses to consider infinitely small variations in a discontinuous way, i.e. in the form of infinitely small jumps of an infinite activity Lévy process. The second agent observes the small variations in a continuous way, i.e. coming from a Brownian motion. Hence the difference between both market models determines a type of model risk.

In the sequel, we intend by robustness or stability of the model, the convergence results when  $\varepsilon$  goes to zero of  $(S_t^{\varepsilon})_{t \in [0,T]}$  and of its related pricing and hedging formulas.

In this paper we focus mainly on the RM strategies under equivalent martingale measures. In particular we consider some specific martingale measures which are commonly used in finance and in electricity markets: the Esscher transform, the minimal entropy martingale measure and the minimal martingale measure. We prove some common properties for these martingale measures in the exponential Lévy setting in addition to those shown in Benth et al. (2013) and Daveloose et al. (2014). Further, we show that under some conditions on the parameters of the stock price process and of the martingale measure, the value of the portfolio, the amount of wealth, the cost and gain process in a RM strategy are robust to the choice of the model. Moreover, we prove the stability of the overall risk of the RM strategy at time zero. The amount of wealth and the gain process in a MVH strategy coincide with those in the RM strategy and hence the convergence results immediately follow. When we assume a fixed initial portfolio value to set up a MVH strategy we obtain a convergence rate for the loss at maturity.

The BSDEJ approach does not provide a robustness result for the *optimal number* of risky assets in a RM strategy as well as in a MVH strategy. In Daveloose et al. (2014) convergence rates for those optimal numbers and the other quantities are computed using Fourier transform techniques.

The paper is organised as follows: in Section 2 we introduce the notations, define the two martingale models for the stock price, and derive the corresponding BSDEJs for the value of



the discounted hedging portfolio. In Section 3 we study the stability of the quadratic hedging strategies towards the choice of the model and obtain convergence rates. In Section 4 we conclude.

# 2. QUADRATIC HEDGING STRATEGIES IN A MARTINGALE SETTING FOR TWO GEOMETRIC LÉVY STOCK PRICE MODELS

Assume a finite time horizon T > 0. The first considered stock price process is determined by the process  $L = (L_t)_{t \in [0,T]}$  which denotes a Lévy process in the filtered complete probability space  $(\Omega, \mathcal{F}, \mathbb{F}, \mathbb{P})$  satisfying the usual hypotheses as defined in Protter (2005). We work with the càdlàg version of the given Lévy process. The characteristic triplet of the Lévy process L is denoted by  $(a, b^2, \ell)$ . We consider a stock price modelled by a geometric Lévy process, i.e. the stock price is given by  $S_t = S_0 e^{L_t}$ ,  $\forall t \in [0, T]$ , where  $S_0 > 0$ . Let t > 0 be the risk-free instantaneous interest rate. The value of the corresponding riskless asset equals  $e^{rt}$ for any time  $t \in [0, T]$ . We denote the discounted stock price process by  $\hat{S}$ . Hence at any time  $t \in [0, T]$  it equals

$$\hat{S}_t = e^{-rt} S_t = S_0 e^{-rt} e^{L_t}.$$

It holds that

(2.1) 
$$d\hat{S}_t = \hat{S}_t \hat{a} dt + \hat{S}_t b dW_t + \hat{S}_t \int_{\mathbb{R}_0} (e^z - 1) \widetilde{N}(dt, dz),$$

where W is a standard Brownian motion independent of the compensated jump measure  $\widetilde{N}$  and

$$\hat{a} = a - r + \frac{1}{2}b^2 + \int_{\mathbb{R}_0} (e^z - 1 - z \mathbb{1}_{\{|z| < 1\}}) \ell(dz).$$

We assume that  $\hat{a} \neq 0$  so that  $\hat{S}$  is not a  $\mathbb{P}$ -martingale. Furthermore  $\hat{S}$  is not deterministic and arbitrage opportunities are excluded (cfr. Tankov (2010)).

In this paper we are interested in the study of quadratic hedging strategies in the martingale setting. Thus we assume the stock price process is observed under a martingale measure  $\widetilde{\mathbb{P}}$  which is equivalent to the historical measure  $\mathbb{P}$ . We consider martingale measures that belong to the class of structure preserving martingale measures, see Jacod and Shiryaev (2002). In this case, the Lévy triplet of the driving process L under  $\widetilde{\mathbb{P}}$  is denoted by  $(\tilde{a}, b^2, \tilde{\ell})$ . Theorem 3.24 in Jacod and Shiryaev (2002) states conditions which are equivalent to the existence of a parameter  $\Theta \in \mathbb{R}$  and a function  $\rho(z; \Theta)$ ,  $z \in \mathbb{R}$ , such that

(2.2) 
$$\int_{\{|z|<1\}} |z(\rho(z;\Theta)-1)|\ell(dz) < \infty,$$

and such that

(2.3) 
$$\tilde{a} = a + b^2 \Theta + \int_{\{|z| < 1\}} z(\rho(z; \Theta) - 1)\ell(dz) \quad \text{and} \quad \tilde{\ell}(dz) = \rho(z; \Theta)\ell(dz).$$

For  $\hat{S}$  to be a martingale under  $\widetilde{\mathbb{P}}$ , the parameter  $\Theta$  should guarantee the following equation

(2.4) 
$$\hat{a}_0 = \tilde{a} - r + \frac{1}{2}b^2 + \int_{\mathbb{R}_0} (e^z - 1 - z \mathbb{1}_{\{|z| < 1\}}) \tilde{\ell}(dz) = 0,$$

From now on we denote the solution of equation (2.4) –when it exists– by  $\Theta_0$  and the equivalent martingale measure by  $\widetilde{\mathbb{P}}_{\Theta_0}$ . Notice that we obtain different martingale measures



 $\widetilde{\mathbb{P}}_{\Theta_0}$  for different choices of the function  $\rho(.;\Theta_0)$ . In the next section we present some known martingale measures for specific functions  $\rho(.;\Theta_0)$  and specific parameters  $\Theta_0$  which solve (2.4).

The relation between the original measure  $\mathbb{P}$  and the martingale measure  $\widetilde{\mathbb{P}}_{\Theta_0}$  is given by

$$\frac{d\widetilde{\mathbb{P}}_{\Theta_0}}{d\mathbb{P}}\Big|_{\mathcal{F}_t} = \exp\Big(b\Theta_0 W_t - \frac{1}{2}b^2\Theta_0^2 t + \int_0^t \int_{\mathbb{R}_0} \log(\rho(z;\Theta_0))\widetilde{N}(ds,dz) + t \int_{\mathbb{R}_0} \Big(\log(\rho(z;\Theta_0)) + 1 - \rho(z;\Theta_0)\Big)\ell(dz)\Big).$$

From the Girsanov theorem (see e.g. Theorem 1.33 in Øksendal and Sulem (2009)) we know that the processes  $W^{\Theta_0}$  and  $\widetilde{N}^{\Theta_0}$  defined by

(2.5) 
$$dW_t^{\Theta_0} = dW_t - b\Theta_0 dt,$$
$$\widetilde{N}^{\Theta_0}(dt, dz) = N(dt, dz) - \rho(z; \Theta_0)\ell(dz)dt = \widetilde{N}(dt, dz) + (1 - \rho(z; \Theta_0))\ell(dz)dt,$$

for all  $t \in [0,T]$  and  $z \in \mathbb{R}_0$ , are a standard Brownian motion and a compensated jump measure under  $\widetilde{\mathbb{P}}_{\Theta_0}$ . Moreover we can rewrite (2.1) as

(2.6) 
$$d\hat{S}_t = \hat{S}_t b dW_t^{\Theta_0} + \hat{S}_t \int_{\mathbb{R}_0} (e^z - 1) \widetilde{N}^{\Theta_0}(dt, dz).$$

We consider a  $\mathcal{F}_T$ -measurable and square integrable random variable  $H_T$  which denotes the payoff of a contract. The discounted payoff equals  $\hat{H}_T = \mathrm{e}^{-rT} H_T$ . In case the discounted stock price process is a martingale, both quadratic hedging strategies, the mean-variance hedging (MVH) and the risk-minimisation (RM) are related to the Galtchouk-Kunita-Watanabe (GKW) decomposition, see Föllmer and Sondermann (1986). In the following we recall the GKW-decomposition of the  $\mathcal{F}_T$ -measurable and square integrable random variable  $\hat{H}_T$  under the martingale measure  $\widetilde{\mathbb{P}}_{\Theta_0}$ 

(2.7) 
$$\hat{H}_T = \widetilde{\mathbb{E}}^{\Theta_0}[\hat{H}_T] + \int_0^T \xi_s^{\Theta_0} d\hat{S}_s + \mathcal{L}_T^{\Theta_0},$$

where  $\widetilde{\mathbb{E}}^{\Theta_0}$  denotes the expectation under  $\widetilde{\mathbb{P}}_{\Theta_0}$ ,  $\xi^{\Theta_0}$  is a predictable process for which we can determine the stochastic integral with respect to  $\hat{S}$ , and  $\mathcal{L}^{\Theta_0}$  is a square integrable  $\widetilde{\mathbb{P}}_{\Theta_0}$ -martingale with  $\mathcal{L}_0^{\Theta_0} = 0$ , such that  $\mathcal{L}^{\Theta_0}$  is  $\widetilde{\mathbb{P}}_{\Theta_0}$ -orthogonal to  $\hat{S}$ .

The quadratic hedging strategies are determined by the process  $\xi^{\Theta_0}$ . It indicates the number of discounted risky assets to hold in the portfolio. The amount invested in the riskless asset is different in both strategies and is determined by the self-financing property for the MVH strategy and by the replicating condition for the RM strategy. See Schweizer (2001) for more details.

We define the process

$$\hat{V}_t^{\Theta_0} = \widetilde{\mathbb{E}}^{\Theta_0}[\hat{H}_T | \mathcal{F}_t], \quad \forall t \in [0, T],$$

which equals the value of the discounted portfolio for the RM strategy. The GKW-decomposition (2.7) implies that

(2.8) 
$$\hat{V}_{t}^{\Theta_{0}} = \hat{V}_{0}^{\Theta_{0}} + \int_{0}^{t} \xi_{s}^{\Theta_{0}} d\hat{S}_{s} + \mathcal{L}_{t}^{\Theta_{0}}, \quad \forall t \in [0, T].$$



Moreover since  $\mathcal{L}^{\Theta_0}$  is a  $\widetilde{\mathbb{P}}_{\Theta_0}$ -martingale, there exist processes  $X^{\Theta_0}$  and  $Y^{\Theta_0}(z)$  such that

(2.9) 
$$\mathcal{L}_{t}^{\Theta_{0}} = \int_{0}^{t} X_{s}^{\Theta_{0}} dW_{s}^{\Theta_{0}} + \int_{0}^{t} \int_{\mathbb{R}_{0}} Y_{s}^{\Theta_{0}}(z) \widetilde{N}^{\Theta_{0}}(ds, dz), \quad \forall t \in [0, T],$$

and which by the  $\widetilde{\mathbb{P}}_{\Theta_0}$ -orthogonality of  $\mathcal{L}^{\Theta_0}$  and  $\hat{S}$  satisfy

(2.10) 
$$X^{\Theta_0}b + \int_{\mathbb{R}_0} Y^{\Theta_0}(z)(e^z - 1)\rho(z;\Theta_0)\ell(dz) = 0.$$

By substituting (2.6) and (2.9) in (2.8), we retrieve

$$d\hat{V}_{t}^{\Theta_{0}} = (\xi_{t}^{\Theta_{0}} \hat{S}_{t} b + X_{t}^{\Theta_{0}}) dW_{t}^{\Theta_{0}} + \int_{\mathbb{R}_{0}} (\xi_{t}^{\Theta_{0}} \hat{S}_{t} (e^{z} - 1) + Y_{t}^{\Theta_{0}}(z)) \widetilde{N}^{\Theta_{0}}(dt, dz).$$

Let  $\hat{\pi}^{\Theta_0} = \xi^{\Theta_0} \hat{S}$  indicate the amount of wealth invested in the discounted risky asset in a quadratic hedging strategy. We conclude that the following BSDEJ holds for the RM strategy

(2.11) 
$$\begin{cases} d\hat{V}_t^{\Theta_0} = A_t^{\Theta_0} dW_t^{\Theta_0} + \int_{\mathbb{R}_0} B_t^{\Theta_0}(z) \widetilde{N}^{\Theta_0}(dt, dz), \\ \hat{V}_T^{\Theta_0} = \hat{H}_T, \end{cases}$$

where

(2.12) 
$$A^{\Theta_0} = \hat{\pi}^{\Theta_0} b + X^{\Theta_0} \text{ and } B^{\Theta_0}(z) = \hat{\pi}^{\Theta_0} (e^z - 1) + Y^{\Theta_0}(z).$$

Since the random variable  $\hat{H}_T$  is square integrable and  $\mathcal{F}_T$ -measurable, we know by Tang and Li (1994) that the BSDEJ (2.11) has a unique solution  $(\hat{V}^{\Theta_0}, A^{\Theta_0}, B^{\Theta_0})$ . This follows from the fact that the drift parameter of  $\hat{V}^{\Theta_0}$  equals zero under  $\widetilde{\mathbb{P}}_{\Theta_0}$  and thus is Lipschitz continuous.

We introduce another Lévy process  $L^{\varepsilon}$ , for  $0 < \varepsilon < 1$ , which is obtained by truncating the jumps of L with absolute size smaller than  $\varepsilon$  and replacing them by an independent Brownian motion which is appropriately scaled. The second stock price process is denoted by  $S^{\varepsilon} = S_0 e^{L^{\varepsilon}}$  and the corresponding discounted stock price process  $\hat{S}^{\varepsilon}$  is thus given by

$$(2.13) d\hat{S}_t^{\varepsilon} = \hat{S}_t^{\varepsilon} \hat{a}_{\varepsilon} dt + \hat{S}_t^{\varepsilon} b dW_t + \hat{S}_t^{\varepsilon} \int_{\{|z| \ge \varepsilon\}} (e^z - 1) \widetilde{N}(dt, dz) + \hat{S}_t^{\varepsilon} G(\varepsilon) d\widetilde{W}_t,$$

for all  $t \in [0,T]$  and  $\hat{S}_0^{\varepsilon} = S_0$ . Herein  $\widetilde{W}$  is a standard Brownian motion independent of L,

(2.14) 
$$G^{2}(\varepsilon) = \int_{\{|z| < \varepsilon\}} (e^{z} - 1)^{2} \ell(dz), \text{ and}$$

$$\hat{a}_{\varepsilon} = a - r + \frac{1}{2} (b^{2} + G^{2}(\varepsilon)) + \int_{\{|z| > \varepsilon\}} (e^{z} - 1 - z \mathbb{1}_{\{|z| < 1\}}) \ell(dz).$$

From now on, we assume that the filtration  $\mathbb F$  is enlarged with the information of the Brownian motion  $\widetilde W$  and we denote the new filtration by  $\widetilde{\mathbb F}$ . Moreover, we also assume absence of arbitrage in this second model. It is clear that the process  $L^{\varepsilon}$  has the Lévy characteristic triplet  $(a,b^2+G^2(\varepsilon),1_{\{|\cdot|\geq\varepsilon\}}\ell)$  under the measure  $\mathbb P$ .

Let  $\widetilde{\mathbb{P}}_{\varepsilon}$  represent a structure preserving martingale measure for  $\hat{S}^{\varepsilon}$ . The characteristic triplet of the driving process  $L^{\varepsilon}$  w.r.t. this martingale measure is denoted by  $(\tilde{a}_{\varepsilon}, b^2 + G^2(\varepsilon), \tilde{\ell}_{\varepsilon})$ .



From Jacod and Shiryaev (2002) (Theorem 3.24), we know that there exist a parameter  $\Theta \in \mathbb{R}$  and a function  $\rho(z;\Theta)$ ,  $z \in \mathbb{R}$ , under certain conditions, such that

(2.15) 
$$\int_{\{\varepsilon < |z| < 1\}} |z(\rho(z;\Theta) - 1)| \ell(dz) < \infty,$$

and such that

(2.16)

$$\tilde{a}_{\varepsilon} = a + (b^2 + G^2(\varepsilon))\Theta + \int_{\{\varepsilon \le |z| < 1\}} z(\rho(z;\Theta) - 1)\ell(dz) \quad \text{and} \quad \tilde{\ell}_{\varepsilon}(dz) = 1_{\{|z| \ge \varepsilon\}}\rho(z;\Theta)\ell(dz).$$

Let us assume that  $\Theta$  solves the following equation

$$\hat{a}_{\varepsilon} = \tilde{a}_{\varepsilon} - r + \frac{1}{2}(b^2 + G^2(\varepsilon)) + \int_{\mathbb{R}_0} (e^z - 1 - z \mathbb{1}_{\{|z| < 1\}}) \tilde{\ell}_{\varepsilon}(dz) = 0,$$

then  $\hat{S}^{\varepsilon}$  is a martingale under  $\widetilde{\mathbb{P}}$ . From now on we indicate the solution of (2.17) –when it exists– as  $\Theta_{\varepsilon}$  and the martingale measure as  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ .

The relation between the original measure  $\mathbb{P}$  and the martingale measure  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$  is given by

$$\begin{split} \frac{d\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}}{d\mathbb{P}}\Big|_{\widetilde{\mathcal{F}}_{t}} &= \exp\Big(b\Theta_{\varepsilon}W_{t} - \frac{1}{2}b^{2}\Theta_{0}^{2}t + G(\varepsilon)\Theta_{\varepsilon}\widetilde{W}_{t} - \frac{1}{2}G^{2}(\varepsilon)\Theta_{\varepsilon}^{2}t + \int_{0}^{t}\int_{\{|z| \geq \varepsilon\}} \log(\rho(z;\Theta_{\varepsilon}))\widetilde{N}(ds,dz) \\ &+ t\int_{\{|z| \geq \varepsilon\}} \Big(\log(\rho(z;\Theta_{\varepsilon})) + 1 - \rho(z;\Theta_{\varepsilon})\Big)\ell(dz)\Big). \end{split}$$

The processes  $W^{\Theta_{\varepsilon}}$ ,  $\widetilde{W}^{\Theta_{\varepsilon}}$ , and  $\widetilde{N}^{\Theta_{\varepsilon}}$  defined by

$$dW_t^{\Theta_{\varepsilon}} = dW_t - b\Theta_{\varepsilon}dt,$$

(2.18) 
$$d\widetilde{W}_t^{\Theta_{\varepsilon}} = d\widetilde{W}_t - G(\varepsilon)\Theta_{\varepsilon}dt,$$

$$\widetilde{N}^{\Theta_{\varepsilon}}(dt,dz) = N(dt,dz) - \rho(z;\Theta_{\varepsilon})\ell(dz)dt = \widetilde{N}(dt,dz) + (1 - \rho(z;\Theta_{\varepsilon}))\ell(dz)dt$$

for all  $t \in [0,T]$  and  $z \in \{z \in \mathbb{R} : |z| \geq \varepsilon\}$ , are two standard Brownian motions and a compensated jump measure under  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$  (see e.g. Theorem 1.33 in Øksendal and Sulem (2009)). Hence the process  $\hat{S}^{\varepsilon}$  is given by

$$(2.19) d\hat{S}_t^{\varepsilon} = \hat{S}_t^{\varepsilon} b dW_t^{\Theta_{\varepsilon}} + \hat{S}_t^{\varepsilon} \int_{\{|z| \ge \varepsilon\}} (e^z - 1) \widetilde{N}^{\Theta_{\varepsilon}} (dt, dz) + \hat{S}_t^{\varepsilon} G(\varepsilon) d\widetilde{W}_t^{\Theta_{\varepsilon}}.$$

We consider a  $\widetilde{\mathcal{F}}_T$ -measurable and square integrable random variable  $H_T^{\varepsilon}$  which is the payoff of a contract. The discounted payoff is denoted by  $\hat{H}_T^{\varepsilon} = \mathrm{e}^{-rT}H_T^{\varepsilon}$ . The GKW-decomposition of  $\hat{H}_T^{\varepsilon}$  under the martingale measure  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$  equals

(2.20) 
$$\hat{H}_T^{\varepsilon} = \widetilde{\mathbb{E}}^{\Theta_{\varepsilon}} [\hat{H}_T^{\varepsilon}] + \int_0^T \xi_s^{\Theta_{\varepsilon}} d\hat{S}_s^{\varepsilon} + \mathcal{L}_T^{\Theta_{\varepsilon}},$$

where  $\widetilde{\mathbb{E}}^{\Theta_{\varepsilon}}$  is the expectation under  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ ,  $\xi^{\Theta_{\varepsilon}}$  is a predictable process for which we can determine the stochastic integral with respect to  $\hat{S}^{\varepsilon}$ , and  $\mathcal{L}^{\Theta_{\varepsilon}}$  is a square integrable  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ -martingale with  $\mathcal{L}_{0}^{\Theta_{\varepsilon}} = 0$ , such that  $\mathcal{L}^{\Theta_{\varepsilon}}$  is  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ -orthogonal to  $\hat{S}^{\varepsilon}$ .

The value of the discounted portfolio for the RM strategy is defined by

$$\hat{V}_t^{\Theta_{\varepsilon}} = \widetilde{\mathbb{E}}^{\Theta_{\varepsilon}} [\hat{H}_T^{\varepsilon} | \widetilde{\mathcal{F}}_t], \quad \forall t \in [0, T].$$



From the GKW-decomposition (2.20) we have

(2.21) 
$$\hat{V}_{t}^{\Theta_{\varepsilon}} = \hat{V}_{0}^{\Theta_{\varepsilon}} + \int_{0}^{t} \xi_{s}^{\Theta_{\varepsilon}} d\hat{S}_{s}^{\varepsilon} + \mathcal{L}_{t}^{\Theta_{\varepsilon}}, \quad \forall t \in [0, T].$$

Moreover since  $\mathcal{L}^{\Theta_{\varepsilon}}$  is a  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ -martingale, there exist processes  $X^{\Theta_{\varepsilon}}$ ,  $Y^{\Theta_{\varepsilon}}(z)$ , and  $Z^{\Theta_{\varepsilon}}$  such that

$$(2.22) \quad \mathcal{L}_{t}^{\Theta_{\varepsilon}} = \int_{0}^{t} X_{s}^{\Theta_{\varepsilon}} dW_{s}^{\Theta_{\varepsilon}} + \int_{0}^{t} \int_{\{|z| > \varepsilon\}} Y_{s}^{\Theta_{\varepsilon}}(z) \widetilde{N}^{\Theta_{\varepsilon}}(ds, dz) + \int_{0}^{t} Z_{s}^{\Theta_{\varepsilon}} d\widetilde{W}_{s}^{\Theta_{\varepsilon}}, \quad \forall t \in [0, T].$$

The  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$ -orthogonality of  $\mathcal{L}^{\Theta_{\varepsilon}}$  and  $\hat{S}^{\varepsilon}$  implies that

(2.23) 
$$X^{\Theta_{\varepsilon}}b + \int_{\{|z| \ge \varepsilon\}} Y^{\Theta_{\varepsilon}}(z)(e^{z} - 1)\rho(z; \Theta_{\varepsilon})\ell(dz) + Z^{\Theta_{\varepsilon}}G(\varepsilon) = 0.$$

Combining (2.19) and (2.22) in (2.21), we get

$$d\hat{V}_{t}^{\Theta_{\varepsilon}} = (\xi_{t}^{\Theta_{\varepsilon}} \hat{S}_{t}^{\varepsilon} b + X_{t}^{\Theta_{\varepsilon}}) dW_{t}^{\Theta_{\varepsilon}} + \int_{\{|z| \geq \varepsilon\}} (\xi_{t}^{\Theta_{\varepsilon}} \hat{S}_{t}^{\varepsilon} (e^{z} - 1) + Y_{t}^{\Theta_{\varepsilon}} (z)) \widetilde{N}^{\Theta_{\varepsilon}} (dt, dz) + (\xi_{t}^{\Theta_{\varepsilon}} \hat{S}_{t}^{\varepsilon} G(\varepsilon) + Z_{t}^{\Theta_{\varepsilon}}) d\widetilde{W}_{t}^{\Theta_{\varepsilon}}.$$

Let  $\hat{\pi}^{\Theta_{\varepsilon}} = \xi^{\Theta_{\varepsilon}} \hat{S}^{\varepsilon}$  denote the amount of wealth invested in the discounted risky asset in the quadratic hedging strategy. We conclude that the following BSDEJ holds for the RM strategy

(2.24) 
$$\begin{cases} d\hat{V}_{t}^{\Theta_{\varepsilon}} = A_{t}^{\Theta_{\varepsilon}} dW_{t}^{\Theta_{\varepsilon}} + \int_{\{|z| \geq \varepsilon\}} B_{t}^{\Theta_{\varepsilon}}(z) \widetilde{N}^{\Theta_{\varepsilon}}(dt, dz) + C_{t}^{\Theta_{\varepsilon}} d\widetilde{W}_{t}^{\Theta_{\varepsilon}}, \\ \hat{V}_{T}^{\Theta_{\varepsilon}} = \hat{H}_{T}^{\varepsilon}, \end{cases}$$

where

$$(2.25) \ A^{\Theta_{\varepsilon}} = \hat{\pi}^{\Theta_{\varepsilon}} b + X^{\Theta_{\varepsilon}}, \quad B^{\Theta_{\varepsilon}}(z) = \hat{\pi}^{\Theta_{\varepsilon}} (\mathrm{e}^z - 1) + Y^{\Theta_{\varepsilon}}(z), \quad \text{and} \quad C^{\Theta_{\varepsilon}} = \hat{\pi}^{\Theta_{\varepsilon}} G(\varepsilon) + Z^{\Theta_{\varepsilon}}.$$

Since the random variable  $\hat{H}_T^{\varepsilon}$  is square integrable and  $\widetilde{\mathcal{F}}_T$ -measurable we know by Tang and Li (1994) that the BSDEJ (2.24) has a unique solution  $(\hat{V}^{\Theta_{\varepsilon}}, A^{\Theta_{\varepsilon}}, B^{\Theta_{\varepsilon}}, C^{\Theta_{\varepsilon}})$ . This results from the fact that the drift parameter of  $\hat{V}^{\Theta_{\varepsilon}}$  equals zero under  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$  and thus is Lipschitz continuous.

### 3. Robustness of the quadratic hedging strategies

The aim of this section is to study the stability of the quadratic hedging strategies to the variation of the model, where we consider the two stock price models defined in (2.1) and (2.13). We study the stability of the RM strategy extensively and at the end of this section we come back to the MVH strategy. Since we work in the martingale setting, we first present some specific martingale measures which are commonly used in finance and in electricity markets. Then we discuss some common properties which are fulfilled by these measures. This is the topic of the next subsection.



3.1. Robustness of the martingale measures. Recall from the previous section that the martingale measures  $\widetilde{\mathbb{P}}_{\Theta_0}$  and  $\widetilde{\mathbb{P}}_{\Theta_{\varepsilon}}$  are determined via the functions  $\rho(.;\Theta_0)$ ,  $\rho(.;\Theta_{\varepsilon})$  and the parameters  $\Theta_0$ ,  $\Theta_{\varepsilon}$ , respectively. We present the following assumptions on these characteristics.

**Assumptions 3.1.** For  $\Theta_0$ ,  $\Theta_{\varepsilon}$ ,  $\rho(.;\Theta_0)$ , and  $\rho(.;\Theta_{\varepsilon})$  satisfying (2.2), (2.3), (2.4), (2.15), (2.16), and (2.17) we assume the following, where C denotes a positive constant and  $\Theta \in \{\Theta_0, \Theta_{\varepsilon}\}.$ 

- (i)  $\Theta_0$  and  $\Theta_{\varepsilon}$  exist and are unique.
- (ii) It holds that

$$|\Theta_0 - \Theta_{\varepsilon}| \le C\widetilde{G}^2(\varepsilon),$$

where  $\widetilde{G}(\varepsilon) = \max(G(\varepsilon), \sigma(\varepsilon))$ . Herein  $\sigma(\varepsilon)$  equals the standard deviation of the jumps of L with size smaller than  $\varepsilon$ , i.e.

$$\sigma^2(\varepsilon) = \int_{\{|z| < \varepsilon\}} z^2 \ell(dz).$$

(iii) On the other hand,  $\Theta_{\varepsilon}$  is uniformly bounded in  $\varepsilon$ , i.e.

$$|\Theta_{\varepsilon}| \leq C.$$

(iv) For all z in  $\{|z| < 1\}$  it holds that

$$|\rho(z;\Theta)| \le C.$$

(v) We have

$$\int_{\{|z|\geq 1\}} \rho^4(z;\Theta)\ell(dz) \leq C.$$

(vi) It is guaranteed that

$$\int_{\mathbb{R}_0} (1 - \rho(z; \Theta))^2 \ell(dz) \le C.$$

(vii) It holds for  $k \in \{2, 4\}$  that

$$\int_{\mathbb{R}_0} \left( \rho(z; \Theta_0) - \rho(z; \Theta_{\varepsilon}) \right)^k \ell(dz) \le C \widetilde{G}^{2k}(\varepsilon).$$

Widely used martingale measures in the exponential Lévy setting are the Esscher transform (ET), minimal entropy martingale measure (MEMM), and minimal martingale measure (MMM), which are specified as follows.

• In order to define the ET we assume that

(3.1) 
$$\int_{\{|z|\geq 1\}} e^{\theta z} \ell(dz) < \infty, \qquad \forall \theta \in \mathbb{R}.$$

The Lévy measures under the ET are given in (2.3) and (2.16) where  $\rho(z;\Theta) = e^{\Theta z}$ . The ET for the first model is then determined by the parameter  $\Theta_0$  satisfying (2.4). For the second model the ET corresponds to the solution  $\Theta_{\varepsilon}$  of (2.17). See Gerber and Shiu (1996) for more details.

• Let us impose that

(3.2) 
$$\int_{\{|z|\geq 1\}} e^{\theta(e^z-1)} \ell(dz) < \infty, \qquad \forall \theta \in \mathbb{R},$$

and that  $\rho(z;\Theta) = e^{\Theta(e^z-1)}$  in the Lévy measures. Then the solution  $\Theta_0$  of equation (2.4) determines the MEMM for the first model, and  $\Theta_{\varepsilon}$  being the solution of (2.17) characterises the MEMM for the second model. The MEMM is studied by Fujiwara and Miyahara (2003).

• Let us consider the assumption

$$(3.3) \qquad \int_{\{z\geq 1\}} e^{4z} \ell(dz) < \infty.$$

The MMM implies that  $\rho(z;\Theta) = \Theta(e^z - 1) - 1$  in the Lévy measures and the parameters  $\Theta_0$  and  $\Theta_{\varepsilon}$  are the solutions of (2.4) and (2.17). More information about the MMM can be found in Arai (2004) and in Föllmer and Schweizer (1991).

In Benth et al. (2013) and Daveloose et al. (2014) it was shown that the ET, the MEMM, and the MMM fulfill statements (i), (ii), (iii), and (iv) of Assumptions 3.1 in the exponential Lévy setting. The following proposition shows that items (v), (vi), and (vii) of Assumptions 3.1 also hold for these martingale measures.

**Proposition 3.2.** The Lévy measures given in (2.3) and (2.16) and corresponding to the ET, MEMM, and MMM, satisfy (v), (vi), and (vii) of Assumptions 3.1.

*Proof.* Recall that the Lévy measure satisfies the following integrability conditions

(3.4) 
$$\int_{\{|z|<1\}} z^2 \ell(dz) < \infty \quad \text{and} \quad \int_{\{|z|\geq 1\}} \ell(dz) < \infty.$$

We show that the statement holds for the considered martingale measures.

• Under the ET it holds for  $\Theta \in \{\Theta_0, \Theta_{\varepsilon}\}$  that

$$\rho^4(z;\Theta) = e^{4\Theta z} \le e^{4C|z|},$$

because of (iii) in Assumptions 3.1. By the mean value theorem (MVT), there exists a number  $\Theta'$  between 0 and  $\Theta$  such that

$$(1 - \rho(z; \Theta))^2 = z^2 e^{2\Theta'z} \Theta^2 \le (1_{\{|z| < 1\}} e^{2C} z^2 + 1_{\{|z| \ge 1\}} e^{(2C+2)z}) C,$$

where we used again Assumptions 3.1 (iii). For  $k \in \{2,4\}$ , we derive via the MVT that

$$(\rho(z;\Theta_0) - \rho(z;\Theta_\varepsilon))^k = e^{k\Theta_0 z} (1 - e^{(\Theta_\varepsilon - \Theta_0)z})^k = e^{k\Theta_0 z} z^k e^{k\Theta'' z} (\Theta_0 - \Theta_\varepsilon)^k,$$

where  $\Theta''$  is a number between 0 and  $\Theta_{\varepsilon} - \Theta_0$ . Assumptions 3.1 (ii) implies that

$$(\rho(z;\Theta_0) - \rho(z;\Theta_\varepsilon))^k \le \left(1_{\{|z|<1\}} e^{k(|\Theta_0|+C)} z^2 + 1_{\{|z|\ge1\}} e^{k(\Theta_0+1+C)z}\right) C\widetilde{G}^{2k}(\varepsilon).$$

The obtained inequalities and integrability conditions (3.1) and (3.4) prove the statement.



• Consider the MEMM and  $\Theta \in \{\Theta_0, \Theta_{\varepsilon}\}$ . We have

$$\rho^4(z;\Theta) = e^{4\Theta(e^z - 1)} \le e^{4C|e^z - 1|},$$

because of (iii) in Assumptions 3.1. The latter assumption and the MVT imply that  $(1 - \rho(z; \Theta))^2 = (e^z - 1)^2 e^{2\Theta'(e^z - 1)}\Theta^2 \le (1_{\{|z| < 1\}} e^{2C(e+1)+2} z^2 + 1_{\{|z| \ge 1\}} e^{(2C+2)(e^z - 1)})C$ .

We determine via the MVT and properties (ii) and (iii) in Assumptions 3.1 for  $k \in \{2,4\}$  that

$$(\rho(z; \Theta_{0}) - \rho(z; \Theta_{\varepsilon}))^{k}$$

$$= e^{k\Theta_{0}(e^{z}-1)} (1 - e^{(\Theta_{\varepsilon}-\Theta_{0})(e^{z}-1)})^{k}$$

$$= e^{k\Theta_{0}(e^{z}-1)} (e^{z} - 1)^{k} e^{k\Theta''(e^{z}-1)} (\Theta_{0} - \Theta_{\varepsilon})^{k}$$

$$\leq \left(1_{\{|z|<1\}} e^{k(|\Theta_{0}|(e+1)+1+C(e+1))} z^{2} + 1_{\{|z|\geq1\}} e^{k(\Theta_{0}+1+C)(e^{z}-1)}\right) C\widetilde{G}^{2k}(\varepsilon).$$

From (3.2) and (3.4) we conclude that (v), (vi), and (vii) in Assumptions 3.1 are in force

• For the MMM we have

$$\rho^4(z;\Theta) = (\Theta(e^z - 1) - 1)^4 \le C(e^{4z} + 1).$$

Moreover it holds that

$$(1 - \rho(z; \Theta))^2 = (e^z - 1)^2 \Theta^2 \le (1_{\{|z| < 1\}} e^z z^2 + 1_{\{|z| \ge 1\}} (e^{zz} + 1)) C.$$

We get through (ii) and (iii) in Assumptions 3.1 that

$$(\rho(z;\Theta_0) - \rho(z;\Theta_{\varepsilon}))^k = (e^z - 1)^k (\Theta_0 - \Theta_{\varepsilon})^k \le \left(1_{\{|z| < 1\}} e^k z^2 + 1_{\{|z| \ge 1\}} (e^{kz} + 1)\right) C\widetilde{G}^{2k}(\varepsilon),$$

for  $k \in \{2,4\}$ . The proof is completed by involving conditions (3.3) and (3.4).

3.2. Robustness of the BSDEJ. In the following lemma we prove the  $L^2$ -boundedness of the solution of the BSDEJ (2.11).

**Lemma 3.3.** Assume point (vi) from Assumptions 3.1. Let  $(\hat{V}^{\Theta_0}, A^{\Theta_0}, B^{\Theta_0})$  be the solution of (2.11). Then we have for all  $t \in [0, T]$ 

$$\mathbb{E}\Big[\int_t^T (\hat{V}_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T (A_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big] \le C \mathbb{E}[\hat{H}_T^2],$$

where C represents a positive constant.

*Proof.* Via (2.5) we rewrite the BSDEJ (2.11) as follows

$$d\hat{V}_t^{\Theta_0} = \left(-b\Theta_0 A_t^{\Theta_0} + \int_{\mathbb{R}_0} B_t^{\Theta_0}(z)(1-\rho(z;\Theta_0))\ell(dz)\right)dt + A_t^{\Theta_0}dW_t + \int_{\mathbb{R}_0} B_t^{\Theta_0}(z)\widetilde{N}(dt,dz).$$

We apply the Itô formula (see e.g. Di Nunno et al. (2009)) to  $e^{\beta t}(\hat{V}_t^{\Theta_0})^2$  and find that

$$d(e^{\beta t}(\hat{V}_{t}^{\Theta_{0}})^{2}) = \beta e^{\beta t}(\hat{V}_{t}^{\Theta_{0}})^{2}dt + 2e^{\beta t}\hat{V}_{t}^{\Theta_{0}}(-b\Theta_{0}A_{t}^{\Theta_{0}} + \int_{\mathbb{R}_{0}} B_{t}^{\Theta_{0}}(z)(1-\rho(z;\Theta_{0}))\ell(dz)dt + 2e^{\beta t}\hat{V}_{t}^{\Theta_{0}}A_{t}^{\Theta_{0}}dW_{t} + e^{\beta t}(A_{t}^{\Theta_{0}})^{2}dt$$



$$+ \int_{\mathbb{R}_0} e^{\beta t} \left( (\hat{V}_{t-}^{\Theta_0} + B_t^{\Theta_0}(z))^2 - (\hat{V}_{t-}^{\Theta_0})^2 \right) \widetilde{N}(dt, dz) + \int_{\mathbb{R}_0} e^{\beta t} (B_t^{\Theta_0}(z))^2 \ell(dz) dt.$$

By integration and taking the expectation we recover that

$$\mathbb{E}[e^{\beta t}(\hat{V}_{t}^{\Theta_{0}})^{2}] = \mathbb{E}[e^{\beta T}(\hat{V}_{T}^{\Theta_{0}})^{2}] - \beta \mathbb{E}\left[\int_{t}^{T} e^{\beta s}(\hat{V}_{s}^{\Theta_{0}})^{2} ds\right] 
- 2\mathbb{E}\left[\int_{t}^{T} e^{\beta s} \hat{V}_{s}^{\Theta_{0}} \left(-b\Theta_{0}A_{s}^{\Theta_{0}} + \int_{\mathbb{R}_{0}} B_{s}^{\Theta_{0}}(z)(1-\rho(z;\Theta_{0}))\ell(dz)\right) ds\right] 
- \mathbb{E}\left[\int_{t}^{T} e^{\beta s}(A_{s}^{\Theta_{0}})^{2} ds\right] - \mathbb{E}\left[\int_{t}^{T} \int_{\mathbb{R}_{0}} e^{\beta s}(B_{s}^{\Theta_{0}}(z))^{2}\ell(dz) ds\right].$$

Because of the properties

(3.6) for all 
$$a, b \in \mathbb{R}$$
 and  $k \in \mathbb{R}_0$  it holds that  $\pm 2ab \le ka^2 + \frac{1}{k}b^2$  and

(3.7) for all  $n \in \mathbb{N}_0$  and for all  $a_i \in \mathbb{R}, i = 1, ..., n$  we have that  $\left(\sum_{i=1}^n a_i\right)^2 \le n \sum_{i=1}^n a_i^2$ ,

the third term in the right hand side of (3.5) is estimated by

$$-2\mathbb{E}\Big[\int_{t}^{T} e^{\beta s} \hat{V}_{s}^{\Theta_{0}} \Big(-b\Theta_{0}A_{s}^{\Theta_{0}} + \int_{\mathbb{R}_{0}} B_{s}^{\Theta_{0}}(z)(1-\rho(z;\Theta_{0}))\ell(dz)\Big)ds\Big]$$

$$\leq \mathbb{E}\Big[\int_{t}^{T} e^{\beta s} \Big\{k(\hat{V}_{s}^{\Theta_{0}})^{2} + \frac{1}{k}\Big(-b\Theta_{0}A_{s}^{\Theta_{0}} + \int_{\mathbb{R}_{0}} B_{s}^{\Theta_{0}}(z)(1-\rho(z;\Theta_{0}))\ell(dz)\Big)^{2}\Big\}ds\Big]$$

$$\leq k\mathbb{E}\Big[\int_{t}^{T} e^{\beta s}(\hat{V}_{s}^{\Theta_{0}})^{2}ds\Big] + \frac{2}{k}b^{2}\Theta_{0}^{2}\mathbb{E}\Big[\int_{t}^{T} e^{\beta s}(A_{s}^{\Theta_{0}})^{2}ds\Big]$$

$$+ \frac{2}{k}\int_{\mathbb{R}_{0}} (1-\rho(z;\Theta_{0}))^{2}\ell(dz)\mathbb{E}\Big[\int_{t}^{T} e^{\beta s}\int_{\mathbb{R}_{0}} (B_{s}^{\Theta_{0}}(z))^{2}\ell(dz)ds\Big].$$

Substituting the latter inequality in (3.5) leads to

$$\mathbb{E}\left[e^{\beta t}(\hat{V}_{t}^{\Theta_{0}})^{2}\right] + (\beta - k)\mathbb{E}\left[\int_{t}^{T}e^{\beta s}(\hat{V}_{s}^{\Theta_{0}})^{2}ds\right] + \left(1 - \frac{2}{k}b^{2}\Theta_{0}^{2}\right)\mathbb{E}\left[\int_{t}^{T}e^{\beta s}(A_{s}^{\Theta_{0}})^{2}ds\right] \\
+ \left(1 - \frac{2}{k}\int_{\mathbb{R}_{0}}\left(1 - \rho(z;\Theta_{0})\right)^{2}\ell(dz)\right)\mathbb{E}\left[\int_{t}^{T}e^{\beta s}\int_{\mathbb{R}_{0}}(B_{s}^{\Theta_{0}}(z))^{2}\ell(dz)ds\right] \\
\leq \mathbb{E}\left[e^{\beta T}(\hat{V}_{T}^{\Theta_{0}})^{2}\right].$$

Let k guarantee that

$$1 - \frac{2}{k}b^2\Theta_0^2 \ge \frac{1}{2}$$
 and  $1 - \frac{2}{k}\int_{\mathbb{R}_0} (1 - \rho(z; \Theta_0))^2 \ell(dz) \ge \frac{1}{2}$ .

Hence we choose

$$k \ge 4 \max \left( b^2 \Theta_0^2, \int_{\mathbb{R}_0} \left( 1 - \rho(z; \Theta_0) \right)^2 \ell(dz) \right) > 0,$$



which exists because of *(vi)* from Assumptions 3.1. Besides we assume that  $\beta \geq k + \frac{1}{2} > 0$ . Then for  $s \in [0, T]$  it follows that  $1 \leq e^{\beta s} \leq e^{\beta T}$  and from (3.8) we achieve

$$\mathbb{E}\Big[\int_t^T (\hat{V}_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T (A_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big] \le C \mathbb{E}[(\hat{V}_T^{\Theta_0})^2],$$

which proves the claim.

The aim of this subsection is to study the robustness of the BSDEJs (2.11) and (2.24). Thereto we consider both models under the enlarged filtration  $\widetilde{\mathbb{F}}$  since we have for all  $t \in [0, T]$  that  $\mathcal{F}_t \subset \widetilde{\mathcal{F}}_t$ . Let us define

$$\bar{V}^{\varepsilon} = \hat{V}^{\Theta_0} - \hat{V}^{\Theta_{\varepsilon}}, \quad \bar{A}^{\varepsilon} = A^{\Theta_0} - A^{\Theta_{\varepsilon}}, \quad \bar{B}^{\varepsilon}(z) = B^{\Theta_0}(z) - 1_{\{|z| \ge \varepsilon\}} B^{\Theta_{\varepsilon}}(z).$$

We derive from (2.5), (2.11), (2.18), and (2.24) that

(3.9) 
$$d\bar{V}_t^{\varepsilon} = \alpha_t^{\varepsilon} dt + \bar{A}_t^{\varepsilon} dW_t + \int_{\mathbb{R}_0} \bar{B}_t^{\varepsilon}(z) \widetilde{N}(dt, dz) - C_t^{\Theta_{\varepsilon}} d\widetilde{W}_t,$$

where

(3.10)

$$\alpha^{\varepsilon} = -b(\Theta_0 A^{\Theta_0} - \Theta_{\varepsilon} A^{\Theta_{\varepsilon}})$$

$$+ \int_{\mathbb{R}_0} \left( B^{\Theta_0}(z) \left( 1 - \rho(z; \Theta_0) \right) - 1_{\{|z| \ge \varepsilon\}} B^{\Theta_{\varepsilon}}(z) \left( 1 - \rho(z; \Theta_{\varepsilon}) \right) \right) \ell(dz) + G(\varepsilon) \Theta_{\varepsilon} C^{\Theta_{\varepsilon}}.$$

The process  $\alpha^{\varepsilon}$  (3.10) plays a crucial role in the study of the robustness of the BSDEJ. In the following lemma we state an upper bound for this process in terms of the solutions of the BSDEJs.

**Lemma 3.4.** Let Assumptions 3.1 hold true. Consider  $\alpha^{\varepsilon}$  as defined in (3.10). For any  $t \in [0,T]$  and  $\beta \in \mathbb{R}$  it holds that

$$\begin{split} \mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} (\alpha_s^\varepsilon)^2 ds\Big] &\leq C\Big(\widetilde{G}^4(\varepsilon) \Big\{\mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} (A_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big]\Big\} \\ &+ \mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} (\bar{A}_s^\varepsilon)^2 ds\Big] + \mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} \int_{\mathbb{R}_0} (\bar{B}_s^\varepsilon(z))^2 \ell(dz) ds\Big] \\ &+ \mathbb{E}\Big[\int_t^T \mathrm{e}^{\beta s} (C_s^{\Theta_\varepsilon})^2 ds\Big]\Big), \end{split}$$

where C is a positive constant.

Proof. Parts (ii) and (iii) of Assumptions 3.1 imply that

$$|-b(\Theta_0 A_s^{\Theta_0} - \Theta_{\varepsilon} A_s^{\Theta_{\varepsilon}})| \le |b||\Theta_0 - \Theta_{\varepsilon}||A_s^{\Theta_0}| + |b||\Theta_{\varepsilon}||A_s^{\Theta_0} - A_s^{\Theta_{\varepsilon}}| \le C\widetilde{G}^2(\varepsilon)|A_s^{\Theta_0}| + C|\bar{A}_s^{\varepsilon}|$$
 and

$$|G(\varepsilon)\Theta_{\varepsilon}C_{s}^{\Theta_{\varepsilon}}| \leq C|C_{s}^{\Theta_{\varepsilon}}|.$$

From Hölder's inequality and Assumptions 3.1 (vi) and (vii) it follows that

$$\left| \int_{\mathbb{R}_{0}} \left( B_{s}^{\Theta_{0}}(z) \left( 1 - \rho(z; \Theta_{0}) \right) - 1_{\{|z| \geq \varepsilon\}} B_{s}^{\Theta_{\varepsilon}}(z) \left( 1 - \rho(z; \Theta_{\varepsilon}) \right) \right) \ell(dz) \right| \\
\leq \left| \int_{\mathbb{R}_{0}} B_{s}^{\Theta_{0}}(z) \left( \rho(z; \Theta_{0}) - \rho(z; \Theta_{\varepsilon}) \right) \ell(dz) \right| + \left| \int_{\mathbb{R}_{0}} \bar{B}_{s}^{\varepsilon}(z) \left( 1 - \rho(z; \Theta_{\varepsilon}) \right) \ell(dz) \right|$$



$$\leq \left(\int_{\mathbb{R}_{0}} \left(\rho(z;\Theta_{0}) - \rho(z;\Theta_{\varepsilon})\right)^{2} \ell(dz)\right)^{1/2} \left(\int_{\mathbb{R}_{0}} (B_{s}^{\Theta_{0}}(z))^{2} \ell(dz)\right)^{1/2} \\
+ \left(\int_{\mathbb{R}_{0}} \left(1 - \rho(z;\Theta_{\varepsilon})\right)^{2} \ell(dz)\right)^{1/2} \left(\int_{\mathbb{R}_{0}} (\bar{B}_{s}^{\varepsilon}(z))^{2} \ell(dz)\right)^{1/2} \\
\leq C\widetilde{G}^{2}(\varepsilon) \left(\int_{\mathbb{R}_{0}} (B_{s}^{\Theta_{0}}(z))^{2} \ell(dz)\right)^{1/2} + C\left(\int_{\mathbb{R}_{0}} (\bar{B}_{s}^{\varepsilon}(z))^{2} \ell(dz)\right)^{1/2}.$$

We conclude that

$$(\alpha_s^{\varepsilon})^2 \le C\Big(\widetilde{G}^4(\varepsilon)\Big\{(A_s^{\Theta_0})^2 + \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz)\Big\} + (\bar{A}_s^{\varepsilon})^2 + \int_{\mathbb{R}_0} (\bar{B}_s^{\varepsilon}(z))^2 \ell(dz) + (C_s^{\Theta_{\varepsilon}})^2\Big).$$

The statement is easily deduced from this inequality.

With these two lemmas ready for use, we state and prove the main result of this subsection which is the robustness of the BSDEJs for the discounted portfolio value process of the RM strategy.

**Theorem 3.5.** Assumptions 3.1 are in force. Let  $(\hat{V}^{\Theta_0}, A^{\Theta_0}, B^{\Theta_0})$  be the solution of (2.11) and  $(\hat{V}^{\Theta_{\varepsilon}}, A^{\Theta_{\varepsilon}}, B^{\Theta_{\varepsilon}}, C^{\Theta_{\varepsilon}})$  be the solution of (2.24). For some positive constant C and any  $t \in [0, T]$  we have

$$\mathbb{E}\Big[\int_{t}^{T} (\hat{V}_{s}^{\Theta_{0}} - \hat{V}_{s}^{\Theta_{\varepsilon}})^{2} ds\Big] + \mathbb{E}\Big[\int_{t}^{T} (A_{s}^{\Theta_{0}} - A_{s}^{\Theta_{\varepsilon}})^{2} ds\Big] \\
+ \mathbb{E}\Big[\int_{t}^{T} \int_{\mathbb{R}_{0}} \left(B_{s}^{\Theta_{0}}(z) - 1_{\{|z| \geq \varepsilon\}} B_{s}^{\Theta_{\varepsilon}}(z)\right)^{2} \ell(dz) ds\Big] + \mathbb{E}\Big[\int_{t}^{T} (C_{s}^{\Theta_{\varepsilon}})^{2} ds\Big] \\
\leq C\Big(\mathbb{E}\Big[(\hat{H}_{T} - \hat{H}_{T}^{\varepsilon})^{2}\Big] + \widetilde{G}^{4}(\varepsilon) \mathbb{E}[\hat{H}_{T}^{2}]\Big).$$

*Proof.* We apply the Itô formula to  $e^{\beta t}(\bar{V}_t^{\varepsilon})^2$ 

$$\begin{split} d \big( \mathrm{e}^{\beta t} (\bar{V}_t^\varepsilon)^2 \big) &= \beta \mathrm{e}^{\beta t} (\bar{V}_t^\varepsilon)^2 dt + 2 \mathrm{e}^{\beta t} \bar{V}_t^\varepsilon \alpha_t^\varepsilon dt + 2 \mathrm{e}^{\beta t} \bar{V}_t^\varepsilon \bar{A}_t^\varepsilon dW_t - 2 \mathrm{e}^{\beta t} \bar{V}_t^\varepsilon C_t^{\Theta_\varepsilon} d\widetilde{W}_t + \mathrm{e}^{\beta t} (\bar{A}_t^\varepsilon)^2 dt \\ &+ \mathrm{e}^{\beta t} (C_t^{\Theta_\varepsilon})^2 dt + \int_{\mathbb{R}_0} \mathrm{e}^{\beta t} \big( (\bar{V}_{t-}^\varepsilon + \bar{B}_t^\varepsilon(z))^2 - (\bar{V}_{t-}^\varepsilon)^2 \big) \widetilde{N}(dt, dz) \\ &+ \int_{\mathbb{R}_0} \mathrm{e}^{\beta t} (\bar{B}_t^\varepsilon(z))^2 \ell(dz) dt. \end{split}$$

Integration over the interval [t,T] and taking the expectation under  $\mathbb P$  results into

$$\mathbb{E}[e^{\beta t}(\bar{V}_{t}^{\varepsilon})^{2}] = \mathbb{E}[e^{\beta T}(\bar{V}_{T}^{\varepsilon})^{2}] - \beta \mathbb{E}\left[\int_{t}^{T} e^{\beta s}(\bar{V}_{s}^{\varepsilon})^{2} ds\right] - 2\mathbb{E}\left[\int_{t}^{T} e^{\beta s}\bar{V}_{s}^{\varepsilon}\alpha_{s}^{\varepsilon} ds\right] - \mathbb{E}\left[\int_{t}^{T} e^{\beta s}(\bar{A}_{s}^{\varepsilon})^{2} ds\right] - \mathbb{E}\left[\int_{t}^{T} \int_{\mathbb{R}_{0}} e^{\beta s}(\bar{B}_{s}^{\varepsilon}(z))^{2} \ell(dz) ds\right] - \mathbb{E}\left[\int_{t}^{T} e^{\beta s}(C_{s}^{\Theta_{\varepsilon}})^{2} ds\right],$$

or equivalently

$$\begin{split} & \mathbb{E}[\mathrm{e}^{\beta t}(\bar{V}_{t}^{\varepsilon})^{2}] + \mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(\bar{A}_{s}^{\varepsilon})^{2}ds\Big] + \mathbb{E}\Big[\int_{t}^{T}\int_{\mathbb{R}_{0}}\mathrm{e}^{\beta s}(\bar{B}_{s}^{\varepsilon}(z))^{2}\ell(dz)ds\Big] + \mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(C_{s}^{\Theta_{\varepsilon}})^{2}ds\Big] \\ & = \mathbb{E}[\mathrm{e}^{\beta T}(\bar{V}_{T}^{\varepsilon})^{2}] - \beta \mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(\bar{V}_{s}^{\varepsilon})^{2}ds\Big] - 2\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}\bar{V}_{s}^{\varepsilon}\alpha_{s}^{\varepsilon}ds\Big] \end{split}$$



(3.11)

$$\leq \mathbb{E}[e^{\beta T}(\bar{V}_T^{\varepsilon})^2] + (k - \beta)\mathbb{E}\Big[\int_t^T e^{\beta s}(\bar{V}_s^{\varepsilon})^2 ds\Big] + \frac{1}{k}\mathbb{E}\Big[\int_t^T e^{\beta s}(\alpha_s^{\varepsilon})^2 ds\Big],$$

where we used property (3.6). The combination of (3.11) with Lemma 3.4 provides

$$\begin{split} \mathbb{E}[\mathrm{e}^{\beta t}(\bar{V}_{t}^{\varepsilon})^{2}] + (\beta - k)\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(\bar{V}_{s}^{\varepsilon})^{2}ds\Big] + \Big(1 - \frac{C}{k}\Big)\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(\bar{A}_{s}^{\varepsilon})^{2}ds\Big] \\ + \Big(1 - \frac{C}{k}\Big)\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}\int_{\mathbb{R}_{0}}(\bar{B}_{s}^{\Theta_{\varepsilon}}(z))^{2}\ell(dz)ds\Big] + \Big(1 - \frac{C}{k}\Big)\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(C_{s}^{\Theta_{\varepsilon}})^{2}ds\Big] \\ (3.12) & \leq \mathbb{E}[\mathrm{e}^{\beta T}(\bar{V}_{T}^{\varepsilon})^{2}] + \frac{C}{k}\widetilde{G}^{4}(\varepsilon)\Big\{\mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}(A_{s}^{\Theta_{0}})^{2}ds\Big] + \mathbb{E}\Big[\int_{t}^{T}\mathrm{e}^{\beta s}\int_{\mathbb{R}_{0}}(B_{s}^{\Theta_{0}}(z))^{2}\ell(dz)ds\Big]\Big\}. \end{split}$$

Let us choose k and  $\beta$  such that  $1 - \frac{C}{k} \ge \frac{1}{2}$  and  $\beta - k \ge \frac{1}{2}$ . This means we choose  $k \ge 2C > 0$  and  $\beta \ge \frac{1}{2} + k > 0$ . Thus for any  $s \in [t, T]$  it holds that  $1 < e^{\beta s} \le e^{\beta T}$ . We derive from (3.12) that

$$\mathbb{E}\Big[\int_{t}^{T} (\bar{V}_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{t}^{T} (\bar{A}_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{t}^{T} \int_{\mathbb{R}_{0}} (\bar{B}_{s}^{\Theta_{\varepsilon}}(z))^{2} \ell(dz) ds\Big] + \mathbb{E}\Big[\int_{t}^{T} (C_{s}^{\Theta_{\varepsilon}})^{2} ds\Big]$$

$$\leq C\Big(\mathbb{E}[(\bar{V}_{T}^{\varepsilon})^{2}] + \widetilde{G}^{4}(\varepsilon)\Big\{\mathbb{E}\Big[\int_{t}^{T} (A_{s}^{\Theta_{0}})^{2} ds\Big] + \mathbb{E}\Big[\int_{t}^{T} \int_{\mathbb{R}_{0}} (B_{s}^{\Theta_{0}}(z))^{2} \ell(dz) ds\Big]\Big\}\Big).$$

By Lemma 3.3 we conclude the proof.

This main result leads to the following theorem concerning the robustness of the discounted portfolio value process of the RM strategy.

**Theorem 3.6.** Assume Assumptions 3.1. Let  $\hat{V}^{\Theta_0}$ ,  $\hat{V}^{\Theta_{\varepsilon}}$  be part of the solution of (2.11), (2.24) respectively. Then we have

$$\mathbb{E}\Big[\sup_{0 < t < T} (\hat{V}_s^{\Theta_0} - \hat{V}_s^{\Theta_{\varepsilon}})^2\Big] \le C\Big(\mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2] + \widetilde{G}^4(\varepsilon)\mathbb{E}[\hat{H}_T^2]\Big),$$

for a positive constant C.

*Proof.* Integration of the BSDEJ (3.9) results into

$$\bar{V}_t^\varepsilon = \bar{V}_T^\varepsilon - \int_t^T \alpha_s^\varepsilon ds - \int_t^T \bar{A}_s^\varepsilon dW_s - \int_t^T \int_{\mathbb{R}_0} \bar{B}_s^\varepsilon(z) \widetilde{N}(ds, dz) + \int_t^T C_s^{\Theta_\varepsilon} d\widetilde{W}_s.$$

By property (3.7) we arrive at

$$\mathbb{E}\Big[\sup_{0\leq t\leq T}(\bar{V}_t^{\varepsilon})^2\Big] \leq 5\Big(\mathbb{E}[(\bar{V}_T^{\varepsilon})^2] + \mathbb{E}\Big[\int_0^T (\alpha_s^{\varepsilon})^2 ds\Big] + \mathbb{E}\Big[\sup_{0\leq t\leq T}\Big(\int_t^T \bar{A}_s^{\varepsilon} dW_s\Big)^2\Big] \\ + \mathbb{E}\Big[\sup_{0\leq t\leq T}\Big(\int_t^T \int_{\mathbb{R}_0} \bar{B}_s^{\varepsilon}(z) \widetilde{N}(ds,dz)\Big)^2\Big] + \mathbb{E}\Big[\sup_{0\leq t\leq T}\Big(\int_t^T C_s^{\Theta_{\varepsilon}} d\widetilde{W}_s\Big)^2\Big]\Big).$$

Burkholder's inequality (see e.g., Tang and Li (1994)) guarantees the existence of a positive constant C such that

$$\mathbb{E}\Big[\sup_{0 \le t \le T} \Big(\int_t^T \bar{A}_s^{\varepsilon} dW_s\Big)^2\Big] \le C \mathbb{E}\Big[\int_0^T (\bar{A}_s^{\varepsilon})^2 ds\Big],$$



$$\mathbb{E}\Big[\sup_{0 \le t \le T} \Big(\int_{t}^{T} \int_{\mathbb{R}_{0}} \bar{B}_{s}^{\varepsilon}(z) \widetilde{N}(ds, dz)\Big)^{2}\Big] \le C \mathbb{E}\Big[\int_{0}^{T} \int_{\mathbb{R}_{0}} (\bar{B}_{s}^{\varepsilon}(z))^{2} \ell(dz) ds\Big],$$

$$\mathbb{E}\Big[\sup_{0 \le t \le T} \Big(\int_{t}^{T} C_{s}^{\Theta_{\varepsilon}} d\widetilde{W}_{s}\Big)^{2}\Big] \le C \mathbb{E}\Big[\int_{0}^{T} (C_{s}^{\Theta_{\varepsilon}})^{2} ds\Big].$$

Applying Lemma 3.4 for  $t = 0, \beta = 0$ , Lemma 3.3, and Theorem 3.5 completes the proof.

3.3. Robustness of the risk-minimising strategy. Theorem 3.6 in the previous subsection concerns the robustness result of the value process of the discounted portfolio in the RM strategy. Before we present the stability of the amount of wealth in the RM strategy, we study the relation between  $\hat{\pi}^{\Theta_0}$  (resp.  $\hat{\pi}^{\Theta_{\varepsilon}}$ ) and the solution of the BSDEJ of type (2.11) (resp. (2.24)) in the first (resp. second) model. Consider the processes  $A^{\Theta_0}$  and  $B^{\Theta_0}(z)$  defined in (2.12), then it holds that

$$\begin{split} A^{\Theta_0}b + \int_{\mathbb{R}_0} B^{\Theta_0}(z) (\mathrm{e}^z - 1) \rho(z; \Theta_0) \ell(dz) \\ &= \hat{\pi}^{\Theta_0} b^2 + X^{\Theta_0} b + \int_{\mathbb{R}_0} \left( \hat{\pi}^{\Theta_0} (\mathrm{e}^z - 1)^2 \rho(z; \Theta_0) + Y^{\Theta_0}(z) (\mathrm{e}^z - 1) \rho(z; \Theta_0) \right) \ell(dz) \\ &= \hat{\pi}^{\Theta_0} \left\{ b^2 + \int_{\mathbb{R}_0} (\mathrm{e}^z - 1)^2 \rho(z; \Theta_0) \ell(dz) \right\} + X^{\Theta_0} b + \int_{\mathbb{R}_0} Y^{\Theta_0}(z) (\mathrm{e}^z - 1) \rho(z; \Theta_0) \ell(dz). \end{split}$$

From property (2.10) we attain that

(3.13) 
$$\hat{\pi}^{\Theta_0} = \frac{1}{\kappa_0} \Big( A^{\Theta_0} b + \int_{\mathbb{R}_0} B^{\Theta_0}(z) (e^z - 1) \rho(z; \Theta_0) \ell(dz) \Big),$$

where  $\kappa_0 = b^2 + \int_{\mathbb{R}_0} (e^z - 1)^2 \rho(z; \Theta_0) \ell(dz)$ . Similarly for the second setting we have for the processes  $A^{\Theta_{\varepsilon}}$ ,  $B^{\Theta_{\varepsilon}}(z)$ , and  $C^{\Theta_{\varepsilon}}$  defined in (2.25) that

$$A^{\Theta_{\varepsilon}}b + \int_{\{|z| \geq \varepsilon\}} B^{\Theta_{\varepsilon}}(z)(e^{z} - 1)\rho(z; \Theta_{\varepsilon})\ell(dz) + C^{\Theta_{\varepsilon}}G(\varepsilon)$$

$$= \hat{\pi}^{\Theta_{\varepsilon}}b^{2} + X^{\Theta_{\varepsilon}}b + \int_{\{|z| \geq \varepsilon\}} \left(\hat{\pi}^{\Theta_{\varepsilon}}(e^{z} - 1)^{2}\rho(z; \Theta_{\varepsilon}) + Y^{\Theta_{\varepsilon}}(z)(e^{z} - 1)\rho(z; \Theta_{\varepsilon})\right)\ell(dz)$$

$$+ \hat{\pi}^{\Theta_{\varepsilon}}G^{2}(\varepsilon) + Z^{\Theta_{\varepsilon}}G(\varepsilon)$$

$$= \hat{\pi}^{\Theta_{\varepsilon}}\left\{b^{2} + \int_{\{|z| \geq \varepsilon\}} (e^{z} - 1)^{2}\rho(z; \Theta_{\varepsilon})\ell(dz) + G^{2}(\varepsilon)\right\}$$

$$+ X^{\Theta_{\varepsilon}}b + \int_{\{|z| \geq \varepsilon\}} Y^{\Theta_{\varepsilon}}(z)(e^{z} - 1)\rho(z; \Theta_{\varepsilon})\ell(dz) + Z^{\Theta_{\varepsilon}}G(\varepsilon).$$

Property (2.23) leads to

(3.14) 
$$\hat{\pi}^{\Theta_{\varepsilon}} = \frac{1}{\kappa_{\varepsilon}} \Big( A^{\Theta_{\varepsilon}} b + \int_{\{|z| \geq \varepsilon\}} B^{\Theta_{\varepsilon}}(z) (e^{z} - 1) \rho(z; \Theta_{\varepsilon}) \ell(dz) + C^{\Theta_{\varepsilon}} G(\varepsilon) \Big),$$
where  $\kappa_{\varepsilon} = b^{2} + \int_{\{|z| \geq \varepsilon\}} (e^{z} - 1)^{2} \rho(z; \Theta_{\varepsilon}) \ell(dz) + G^{2}(\varepsilon).$ 

We introduce the following additional assumption on the Lévy measure which we need for the robustness results studied later.



**Assumption 3.7.** For the Lévy measure  $\ell$  the following integrability condition holds

$$\int_{\{z\geq 1\}} e^{4z} \ell(dz) < \infty.$$

Note that the latter assumption, combined with (3.4), implies for  $k \in \{2,4\}$  that

$$(3.15) \qquad \int_{\mathbb{R}_0} (e^z - 1)^k \ell(dz) \le C \left( \int_{\{|z| < 1\}} z^2 \ell(dz) + \int_{\{|z| \ge 1\}} \ell(dz) + \int_{\{z \ge 1\}} e^{4z} \ell(dz) \right) < \infty.$$

Moreover Assumption 3.7 is fulfilled for the considered martingale measures described in subsection 3.1. Indeed, consider the ET, applying (3.1) for  $\theta = 4$  and restricting the integral over  $\{z \geq 1\}$  implies Assumption 3.7. On the set  $\{z \geq 1\}$  it holds that  $z \leq e^z - 1$  and therefore Assumption 3.7 follows from (3.2) by choosing  $\theta = 4$ . For the MMM, condition (3.3) corresponds exactly to Assumption 3.7.

**Theorem 3.8.** Impose Assumptions 3.1 and 3.7. Let the processes  $\hat{\pi}^{\Theta_0}$  and  $\hat{\pi}^{\Theta_{\varepsilon}}$  denote the amounts of wealth in a RM strategy. There is a positive constant C such that for any  $t \in [0,T]$ 

$$\mathbb{E}\Big[\int_{t}^{T} (\hat{\pi}_{s}^{\Theta_{0}} - \hat{\pi}_{s}^{\Theta_{\varepsilon}})^{2} ds\Big] \leq C\Big(\mathbb{E}[(\hat{H}_{T} - \hat{H}_{T}^{\varepsilon})^{2}] + \widetilde{G}^{4}(\varepsilon)\mathbb{E}[\hat{H}_{T}^{2}]\Big).$$

*Proof.* Consider the amounts of wealth in (3.13) and (3.14). Let us denote  $\hat{\pi}^{\Theta_0} = \frac{1}{\kappa_0} \Upsilon^0$  and  $\hat{\pi}^{\Theta_{\varepsilon}} = \frac{1}{\kappa_{\varepsilon}} \Upsilon^{\varepsilon}$ . Then it holds that

$$(\hat{\pi}^{\Theta_0} - \hat{\pi}^{\Theta_{\varepsilon}})^2 \le 2 \left( \left( \frac{\kappa_0 - \kappa_{\varepsilon}}{\kappa_0 \kappa_{\varepsilon}} \right)^2 (\Upsilon^0)^2 + \frac{1}{\kappa_{\varepsilon}^2} (\Upsilon^0 - \Upsilon^{\varepsilon})^2 \right).$$

Herein we have because of the Holder's inequality, (2.14), (3.15), and properties (iv) and (vii) in Assumptions 3.1 that

$$\left(\frac{\kappa_0 - \kappa_{\varepsilon}}{\kappa_0 \kappa_{\varepsilon}}\right)^2 \leq \frac{3}{b^8} \left( \left( \int_{\{|z| < \varepsilon\}} (e^z - 1)^2 \rho(z; \Theta_0) \ell(dz) \right)^2 + \left( \int_{\{|z| \geq \varepsilon\}} (e^z - 1)^2 (\rho(z; \Theta_0) - \rho(z; \Theta_{\varepsilon})) \ell(dz) \right)^2 + G^4(\varepsilon) \right) \\
\leq \frac{3}{b^8} \left( C \left( \int_{\{|z| < \varepsilon\}} (e^z - 1)^2 \ell(dz) \right)^2 + \int_{\mathbb{R}_0} (e^z - 1)^4 \ell(dz) \int_{\mathbb{R}_0} \left( (\rho(z; \Theta_0) - \rho(z; \Theta_{\varepsilon})) \right)^2 \ell(dz) + G^4(\varepsilon) \right) \leq C\widetilde{G}^4(\varepsilon).$$

On the other hand it is clear from (3.13) and (3.14) that

$$(\Upsilon^{0} - \Upsilon^{\varepsilon})^{2} \leq 3 \Big( (\bar{A}^{\varepsilon})^{2} b^{2} + (C^{\Theta_{\varepsilon}})^{2} G^{2}(\varepsilon) + \Big( \int_{\mathbb{R}_{0}} (B^{\Theta_{0}}(z)(e^{z} - 1)\rho(z; \Theta_{0}) - 1_{\{|z| \geq \varepsilon\}} B^{\Theta_{\varepsilon}}(z)(e^{z} - 1)\rho(z; \Theta_{\varepsilon})) \ell(dz) \Big)^{2} \Big).$$

Herein we derive via Holder's inequality, (3.4), (3.15), and points (iv), (v), and (vii) in Assumptions 3.1 that

$$\left(\int_{\mathbb{R}_0} (B^{\Theta_0}(z)(e^z - 1)\rho(z;\Theta_0) - 1_{\{|z| \ge \varepsilon\}} B^{\Theta_\varepsilon}(z)(e^z - 1)\rho(z;\Theta_\varepsilon))\ell(dz)\right)^2$$



$$\begin{split} &= \Big(\int_{\mathbb{R}_{0}} (B^{\Theta_{0}}(z)(\rho(z;\Theta_{0}) - \rho(z;\Theta_{\varepsilon}))(\mathrm{e}^{z} - 1) + \bar{B}^{\varepsilon}(z)\rho(z;\Theta_{\varepsilon})(\mathrm{e}^{z} - 1))\ell(dz)\Big)^{2} \\ &\leq \int_{\mathbb{R}_{0}} (B^{\Theta_{0}}(z))^{2}\ell(dz) \int_{\mathbb{R}_{0}} (\rho(z;\Theta_{0}) - \rho(z;\Theta_{\varepsilon}))^{2}(\mathrm{e}^{z} - 1)^{2}\ell(dz) \\ &+ \int_{\mathbb{R}_{0}} (\bar{B}^{\varepsilon}(z))^{2}\ell(dz) \int_{\mathbb{R}_{0}} \rho^{2}(z;\Theta_{\varepsilon})(\mathrm{e}^{z} - 1)^{2}\ell(dz) \\ &\leq \int_{\mathbb{R}_{0}} (B^{\Theta_{0}}(z))^{2}\ell(dz) \Big(\int_{\mathbb{R}_{0}} (\rho(z;\Theta_{0}) - \rho(z;\Theta_{\varepsilon}))^{4}\ell(dz)\Big)^{\frac{1}{2}} \Big(\int_{\mathbb{R}_{0}} (\mathrm{e}^{z} - 1)^{4}\ell(dz)\Big)^{\frac{1}{2}} \\ &+ \int_{\mathbb{R}_{0}} (\bar{B}^{\varepsilon}(z))^{2}\ell(dz) \Big\{C\int_{\{|z|<1\}} z^{2}\ell(dz) + \Big(\int_{\{|z|\geq1\}} \rho^{4}(z;\Theta_{\varepsilon})\ell(dz)\int_{\{|z|\geq1\}} (\mathrm{e}^{z} - 1)^{4}\ell(dz)\Big)^{\frac{1}{2}} \Big\} \\ &\leq C\widetilde{G}^{4}(\varepsilon)\int_{\mathbb{R}_{0}} (B^{\Theta_{0}}(z))^{2}\ell(dz) + C\int_{\mathbb{R}_{0}} (\bar{B}^{\varepsilon}(z))^{2}\ell(dz). \end{split}$$

The results above show that

$$(\hat{\pi}_t^{\Theta_0} - \hat{\pi}_t^{\Theta_\varepsilon})^2 \leq C\Big((\bar{A}_t^\varepsilon)^2 + \int_{\mathbb{R}_0} (\bar{B}_t^\varepsilon(z))^2 \ell(dz) + (C_t^{\Theta_\varepsilon})^2 + \widetilde{G}^4(\varepsilon) \Big\{ (A_t^{\Theta_0})^2 + \int_{\mathbb{R}_0} (B_t^{\Theta_0}(z))^2 \ell(dz) \Big\} \Big).$$

Therefore

$$\begin{split} \mathbb{E}\Big[\int_t^T (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_\varepsilon})^2 ds\Big] &\leq C\Big(\mathbb{E}\Big[\int_t^T (\bar{A}_s^\varepsilon)^2 ds\Big] + \mathbb{E}\Big[\int_t^T \int_{\mathbb{R}_0} (\bar{B}_s^\varepsilon(z))^2 \ell(dz) ds\Big] + \mathbb{E}\Big[\int_t^T (C_s^{\Theta_\varepsilon})^2 ds\Big] \\ &\quad + \widetilde{G}^4(\varepsilon)\Big\{\mathbb{E}\Big[\int_t^T (A_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_t^T \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big]\Big\}\Big). \end{split}$$

By Lemma 3.3 and Theorem 3.5 we conclude the proof.

The trading in the risky assets is gathered in the gain processes defined by  $\hat{G}_t^{\Theta_0} = \int_0^t \xi_s^{\Theta_0} d\hat{S}_s$  and  $\hat{G}_t^{\Theta_\varepsilon} = \int_0^t \xi_s^{\Theta_\varepsilon} d\hat{S}_s^{\varepsilon}$ . The following theorem shows the robustness of this gain process.

**Theorem 3.9.** Under Assumptions 3.1 and 3.7, there exists a positive constant C such that for any  $t \in [0, T]$ 

$$E\left[\left(\hat{G}_t^{\Theta_0} - \hat{G}_t^{\Theta_\varepsilon}\right)^2\right] \le C\left(\mathbb{E}\left[\left(\hat{H}_T - \hat{H}_T^\varepsilon\right)^2\right] + \widetilde{G}^2(\varepsilon)\mathbb{E}\left[\hat{H}_T^2\right]\right).$$

*Proof.* From (2.5) and (2.6) we know that

$$\begin{aligned} \xi_s^{\Theta_0} d\hat{S}_s &= \xi_s^{\Theta_0} \hat{S}_s b dW_s^{\Theta_0} + \xi_s^{\Theta_0} \hat{S}_s \int_{\mathbb{R}_0} (\mathbf{e}^z - 1) \widetilde{N}^{\Theta_0} (ds, dz) \\ &= \hat{\pi}_s^{\Theta_0} \Big( \Big( -b^2 \Theta_0 + \int_{\mathbb{R}_0} (\mathbf{e}^z - 1) \Big( 1 - \rho(z; \Theta_0) \Big) \ell(dz) \Big) ds + b dW_s + \int_{\mathbb{R}_0} (\mathbf{e}^z - 1) \widetilde{N} (ds, dz) \Big). \end{aligned}$$

In the other setting we have from (2.18) and (2.19) that

$$\xi_s^{\Theta_{\varepsilon}} d\hat{S}_s^{\varepsilon} = \xi_s^{\Theta_{\varepsilon}} \hat{S}_s^{\varepsilon} b dW_s^{\Theta_{\varepsilon}} + \xi_s^{\Theta_{\varepsilon}} \hat{S}_s^{\varepsilon} \int_{\{|z| \ge \varepsilon\}} (e^z - 1) \widetilde{N}^{\Theta_{\varepsilon}} (ds, dz) + \xi_s^{\Theta_{\varepsilon}} \hat{S}_s^{\varepsilon} G(\varepsilon) d\widetilde{W}_s^{\Theta_{\varepsilon}} 
= \hat{\pi}_s^{\Theta_{\varepsilon}} \Big( \Big( -b^2 \Theta_{\varepsilon} + \int_{\{|z| \ge \varepsilon\}} (e^z - 1) \Big( 1 - \rho(z; \Theta_{\varepsilon}) \Big) \ell(dz) - G^2(\varepsilon) \Theta_{\varepsilon} \Big) ds 
+ b dW_s + \int_{\{|z| \ge \varepsilon\}} (e^z - 1) \widetilde{N} (ds, dz) + G(\varepsilon) d\widetilde{W}_s \Big).$$



We derive from the previous SDEs that

$$\begin{split} \hat{G}_{t}^{\Theta_{0}} - \hat{G}_{t}^{\Theta_{\varepsilon}} &= \int_{0}^{t} \xi_{s}^{\Theta_{0}} d\hat{S}_{s} - \int_{0}^{t} \xi_{s}^{\Theta_{\varepsilon}} d\hat{S}_{s}^{\varepsilon} \\ &= \Big( - b^{2}\Theta_{0} + \int_{\mathbb{R}_{0}} (\mathbf{e}^{z} - 1) \Big( 1 - \rho(z; \Theta_{0}) \Big) \ell(dz) \Big) \int_{0}^{t} \hat{\pi}_{s}^{\Theta_{0}} ds \\ &- \Big( - b^{2}\Theta_{\varepsilon} + \int_{\{|z| \geq \varepsilon\}} (\mathbf{e}^{z} - 1) \Big( 1 - \rho(z; \Theta_{\varepsilon}) \Big) \ell(dz) - G^{2}(\varepsilon) \Theta_{\varepsilon} \Big) \int_{0}^{t} \hat{\pi}_{s}^{\Theta_{\varepsilon}} ds \\ &+ b \int_{0}^{t} (\hat{\pi}_{s}^{\Theta_{0}} - \hat{\pi}_{s}^{\Theta_{\varepsilon}}) dW_{s} + \int_{0}^{t} \int_{\mathbb{R}_{0}} (\hat{\pi}_{s}^{\Theta_{0}} (\mathbf{e}^{z} - 1) - \hat{\pi}_{s}^{\Theta_{\varepsilon}} 1_{\{|z| \geq \varepsilon\}} (\mathbf{e}^{z} - 1) \Big) \widetilde{N}(ds, dz) \\ &- G(\varepsilon) \int_{0}^{t} \hat{\pi}_{s}^{\Theta_{\varepsilon}} d\widetilde{W}_{s}. \end{split}$$

Via the Cauchy-Schwartz inequality and the Itô isometry we obtain that

$$\begin{split} &\mathbb{E}\big[\big(\hat{G}_t^{\Theta_0} - \hat{G}_t^{\Theta_\varepsilon}\big)^2\big] \\ &\leq C\Big(\Big\{\Big(-b^2\Theta_0 + \int_{\mathbb{R}_0} (\mathrm{e}^z - 1)\big(1 - \rho(z;\Theta_0)\big)\ell(dz)\Big) \\ &- \Big(-b^2\Theta_\varepsilon + \int_{\{|z| \geq \varepsilon\}} (\mathrm{e}^z - 1)\big(1 - \rho(z;\Theta_\varepsilon)\big)\ell(dz) - G^2(\varepsilon)\Theta_\varepsilon\Big)\Big\}^2 \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big] \\ &+ \Big(-b^2\Theta_\varepsilon + \int_{\{|z| \geq \varepsilon\}} (\mathrm{e}^z - 1)\big(1 - \rho(z;\Theta_\varepsilon)\big)\ell(dz) - G^2(\varepsilon)\Theta_\varepsilon\Big)^2 \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_\varepsilon})^2 ds\Big] \\ &+ b^2 \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_\varepsilon})^2 ds\Big] + \mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (\hat{\pi}_s^{\Theta_0} (\mathrm{e}^z - 1) - \hat{\pi}_s^{\Theta_\varepsilon} 1_{\{|z| \geq \varepsilon\}} (\mathrm{e}^z - 1)\big)^2 \ell(dz) ds\Big] \\ &+ G^2(\varepsilon) \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_\varepsilon})^2 ds\Big]\Big), \end{split}$$

wherein

$$\begin{split} & \mathbb{E}\Big[\int_{0}^{t} \int_{\mathbb{R}_{0}} \left(\hat{\pi}_{s}^{\Theta_{0}}(\mathbf{e}^{z}-1) - \hat{\pi}_{s}^{\Theta_{\varepsilon}} \mathbf{1}_{\{|z| \geq \varepsilon\}}(\mathbf{e}^{z}-1)\right)^{2} \ell(dz) ds \Big] \\ & \leq 2\mathbb{E}\Big[\int_{0}^{t} \int_{\mathbb{R}_{0}} \left((\hat{\pi}_{s}^{\Theta_{0}})^{2}(\mathbf{e}^{z}-1)^{2} \mathbf{1}_{\{|z| < \varepsilon\}} + (\hat{\pi}_{s}^{\Theta_{0}} - \hat{\pi}_{s}^{\Theta_{\varepsilon}})^{2}(\mathbf{e}^{z}-1)^{2} \mathbf{1}_{\{|z| \geq \varepsilon\}}\right) \ell(dz) ds \Big] \\ & \leq 2\Big(\int_{\{|z| < \varepsilon\}} (\mathbf{e}^{z}-1)^{2} \ell(dz) \mathbb{E}\Big[\int_{0}^{t} (\hat{\pi}_{s}^{\Theta_{0}})^{2} ds\Big] + \int_{\mathbb{R}_{0}} (\mathbf{e}^{z}-1)^{2} \ell(dz) \mathbb{E}\Big[\int_{0}^{t} (\hat{\pi}_{s}^{\Theta_{0}} - \hat{\pi}_{s}^{\Theta_{\varepsilon}})^{2} ds\Big]\Big), \end{split}$$

and

$$\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_{\varepsilon}})^2 ds\Big] \le 2\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_{\varepsilon}})^2 ds\Big] + 2\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big].$$

By (2.14), Assumptions 3.1, (3.13), (3.15), Lemma 3.3, and Theorem 3.8 we prove the statement.

The following result shows the robustness of the process  $\mathcal{L}^{\Theta}$  appearing in the GKW-decomposition. This plays an important role in the stability of the cost process of the RM strategy.



**Theorem 3.10.** Let Assumptions 3.1 and 3.7 hold true. Let the processes  $\mathcal{L}^{\Theta_0}$  and  $\mathcal{L}^{\Theta_{\varepsilon}}$  be as in (2.9) and (2.22), respectively. For any  $t \in [0,T]$  it holds that

$$\mathbb{E}[(\mathcal{L}_t^{\Theta_0} - \mathcal{L}_t^{\Theta_{\varepsilon}})^2] \le C\Big(\mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2] + \widetilde{G}^2(\varepsilon)\mathbb{E}[\hat{H}_T^2]\Big).$$

for a positive constant C.

*Proof.* By (2.5) we can rewrite (2.9) as

$$d\mathcal{L}_{t}^{\Theta_{0}} = \left(-b\Theta_{0}X_{t}^{\Theta_{0}} + \int_{\mathbb{R}_{0}}Y_{t}^{\Theta_{0}}(z)(1-\rho(z;\Theta_{0}))\ell(dz)\right)dt + X_{t}^{\Theta_{0}}dW_{t} + \int_{\mathbb{R}_{0}}Y_{t}^{\Theta_{0}}(z)\widetilde{N}(dt,dz).$$

and similarly by (2.18) we obtain for (2.22)

$$d\mathcal{L}_{t}^{\Theta_{\varepsilon}} = \left(-b\Theta_{\varepsilon}X_{t}^{\Theta_{\varepsilon}} + \int_{\{|z| \geq \varepsilon\}} Y_{t}^{\Theta_{\varepsilon}}(z)(1 - \rho(z; \Theta_{\varepsilon}))\ell(dz) - G(\varepsilon)\Theta_{\varepsilon}Z_{t}^{\Theta_{\varepsilon}}\right)dt$$
$$+ X_{t}^{\Theta_{\varepsilon}}dW_{t} + \int_{\{|z| \geq \varepsilon\}} Y_{t}^{\Theta_{\varepsilon}}(z)\widetilde{N}(dt, dz) + Z_{t}^{\Theta_{\varepsilon}}d\widetilde{W}_{t}.$$

Hence we recover that

$$d(\mathcal{L}_{t}^{\Theta_{0}} - \mathcal{L}_{t}^{\Theta_{\varepsilon}}) = \gamma_{t}^{\varepsilon} dt + \bar{X}_{t}^{\varepsilon} dW_{t} + \int_{\mathbb{R}_{0}} \bar{Y}_{t}^{\varepsilon}(z) \widetilde{N}(dt, dz) - Z_{t}^{\Theta_{\varepsilon}} d\widetilde{W}_{t},$$

where

$$\begin{split} \gamma^{\varepsilon} &= -b(\Theta_{0}X^{\Theta_{0}} - \Theta_{\varepsilon}X^{\Theta_{\varepsilon}}) + G(\varepsilon)\Theta_{\varepsilon}Z^{\Theta_{\varepsilon}} \\ &+ \int_{\mathbb{R}_{0}} \left(Y^{\Theta_{0}}(z)(1 - \rho(z;\Theta_{0})) - 1_{\{|z| \geq \varepsilon\}}Y^{\Theta_{\varepsilon}}(z)(1 - \rho(z;\Theta_{\varepsilon}))\right)\ell(dz), \\ \bar{X}^{\varepsilon} &= X^{\Theta_{0}} - X^{\Theta_{\varepsilon}}, \\ \bar{Y}^{\varepsilon}(z) &= Y^{\Theta_{0}}(z) - 1_{\{|z| \geq \varepsilon\}}Y^{\Theta_{\varepsilon}}(z). \end{split}$$

By integration over [0,t] and taking the square we retrieve using (3.7) that

$$(\mathcal{L}_t^{\Theta_0} - \mathcal{L}_t^{\Theta_{\varepsilon}})^2$$

$$\leq C \Big( \Big( \int_0^t \gamma_s^\varepsilon ds \Big)^2 + \Big( \int_0^t \bar{X}_s^\varepsilon dW_s \Big)^2 + \Big( \int_0^t \int_{\mathbb{R}_0} \bar{Y}_s^\varepsilon (z) \widetilde{N}(ds,dz) \Big)^2 + \Big( \int_0^t Z_s^{\Theta_\varepsilon} d\widetilde{W}_s \Big)^2 \Big).$$

Via the Cauchy-Schwartz inequality and the Itô isometry it follows that

$$\mathbb{E}[(\mathcal{L}_{t}^{\Theta_{0}} - \mathcal{L}_{t}^{\Theta_{\varepsilon}})^{2}] \leq C\Big(\mathbb{E}\Big[\int_{0}^{t} (\gamma_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} (\bar{X}_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} (\bar{X}_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} (Z_{s}^{\Theta_{\varepsilon}})^{2} ds\Big]\Big).$$

Concerning the term  $\mathbb{E}\left[\int_0^t (\gamma_s^{\varepsilon})^2 ds\right]$  we derive through *(ii)* and *(iii)* in Assumptions 3.1 that

$$\mathbb{E}\Big[\int_{0}^{t} (\Theta_{0}X_{s}^{\Theta_{0}} - \Theta_{\varepsilon}X_{s}^{\Theta_{\varepsilon}})^{2} ds\Big] \leq 2\Big(\mathbb{E}\Big[\int_{0}^{t} (\Theta_{0} - \Theta_{\varepsilon})^{2} (X_{s}^{\Theta_{0}})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} \Theta_{\varepsilon}^{2} (X_{s}^{\Theta_{0}} - X_{s}^{\Theta_{\varepsilon}})^{2} ds\Big]\Big) \\
\leq C\Big(\widetilde{G}^{4}(\varepsilon)\mathbb{E}\Big[\int_{0}^{t} (X_{s}^{\Theta_{0}})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} (\bar{X}_{s}^{\varepsilon})^{2} ds\Big]\Big)$$



and via (vi) and (vii) in Assumptions 3.1 it follows that

$$\mathbb{E}\Big[\int_{0}^{t} \Big\{ \int_{\mathbb{R}_{0}} \Big( Y_{s}^{\Theta_{0}}(z) \Big( 1 - \rho(z; \Theta_{0}) \Big) - 1_{\{|z| \geq \varepsilon\}} Y_{s}^{\Theta_{\varepsilon}}(z) \Big( 1 - \rho(z; \Theta_{\varepsilon}) \Big) \Big) \ell(dz) \Big\}^{2} ds \Big] \\
\leq \int_{\mathbb{R}_{0}} \Big( \rho(z; \Theta_{0}) - \rho(z; \Theta_{\varepsilon}) \Big)^{2} \ell(dz) \mathbb{E}\Big[ \int_{0}^{t} \int_{\mathbb{R}_{0}} (Y_{s}^{\Theta_{0}}(z))^{2} \ell(dz) ds \Big] \\
+ \int_{\mathbb{R}_{0}} \Big( 1 - \rho(z; \Theta_{\varepsilon}) \Big)^{2} \ell(dz) \mathbb{E}\Big[ \int_{0}^{t} \int_{\mathbb{R}_{0}} (\bar{Y}_{s}^{\varepsilon}(z))^{2} \ell(dz) ds \Big] \\
\leq C\Big( \widetilde{G}^{4}(\varepsilon) \mathbb{E}\Big[ \int_{0}^{t} \int_{\mathbb{R}_{0}} (Y_{s}^{\Theta_{0}}(z))^{2} \ell(dz) ds \Big] + \mathbb{E}\Big[ \int_{0}^{t} \int_{\mathbb{R}_{0}} (\bar{Y}_{s}^{\varepsilon}(z))^{2} \ell(dz) ds \Big] \Big).$$

Thus we obtain that

$$\mathbb{E}[(\mathcal{L}_{t}^{\Theta_{0}} - \mathcal{L}_{t}^{\Theta_{\varepsilon}})^{2}] \leq C\Big(\widetilde{G}^{4}(\varepsilon)\Big\{\mathbb{E}\Big[\int_{0}^{t} (X_{s}^{\Theta_{0}})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} \int_{\mathbb{R}_{0}} (Y_{s}^{\Theta_{0}}(z))^{2} \ell(dz) ds\Big]\Big\} \\
+ \mathbb{E}\Big[\int_{0}^{t} (\bar{X}_{s}^{\varepsilon})^{2} ds\Big] + \mathbb{E}\Big[\int_{0}^{t} \int_{\mathbb{R}_{0}} (\bar{Y}_{s}^{\varepsilon}(z))^{2} \ell(dz) ds\Big] + \mathbb{E}\Big[\int_{0}^{t} (Z_{s}^{\Theta_{\varepsilon}})^{2} ds\Big]\Big).$$

Let us consider the terms appearing in the latter expression separately.

• Definition (2.12) implies that

$$\mathbb{E}\Big[\int_0^t (X_s^{\Theta_0})^2 ds\Big] \le 2\Big(\mathbb{E}\Big[\int_0^t (A_s^{\Theta_0})^2 ds\Big] + b^2 \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big]\Big)$$

and

$$\mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (Y_s^{\Theta_0}(z))^2 \ell(dz) ds\Big] \\
\leq 2\Big(\mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big] + \int_{\mathbb{R}_0} (e^z - 1)^2 \ell(dz) \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big]\Big).$$

• Combining (2.12) and (2.25) in

$$\bar{X}_t^{\varepsilon} = X_t^{\Theta_0} - X_t^{\Theta_{\varepsilon}} = \bar{A}_t^{\varepsilon} - (\hat{\pi}_t^{\Theta_0} - \hat{\pi}_t^{\Theta_{\varepsilon}})b,$$

it easily follows that

$$\mathbb{E}\Big[\int_0^t (\bar{X}_s^{\varepsilon})^2 ds\Big] \le C\Big(\mathbb{E}\Big[\int_0^t (\bar{A}_s^{\varepsilon})^2 ds\Big] + \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_{\varepsilon}})^2 ds\Big]\Big).$$

 $\bullet$  Similarly, from (2.12) and (2.25) we find

$$\bar{Y}_t^{\varepsilon}(z) = Y_t^{\Theta_0}(z) - Y_t^{\Theta_{\varepsilon}}(z) = \bar{B}_t^{\varepsilon}(z) - (\hat{\pi}_t^{\Theta_0} - \hat{\pi}_t^{\Theta_{\varepsilon}})(e^z - 1).$$

Hence

$$\mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (\bar{Y}_s^{\varepsilon}(z))^2 \ell(dz) ds\Big] \\
\leq 2\Big(\mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (\bar{B}_s^{\varepsilon}(z))^2 \ell(dz) ds\Big] + \int_{\mathbb{R}_0} (e^z - 1)^2 \ell(dz) \mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_{\varepsilon}})^2 ds\Big]\Big).$$



• From (2.25), the estimate

$$(Z_t^{\Theta_{\varepsilon}}(z))^2 \le C \left( (C_t^{\Theta_{\varepsilon}})^2 + (\hat{\pi}_t^{\Theta_0} - \hat{\pi}_t^{\Theta_{\varepsilon}})^2 G^2(\varepsilon) + (\hat{\pi}_t^{\Theta_0})^2 G^2(\varepsilon) \right)$$

leads to

$$\mathbb{E}\Big[\int_0^t (Z_s^{\Theta_{\varepsilon}}(z))^2 ds\Big] \\
\leq C\Big(\mathbb{E}\Big[\int_0^t (C_s^{\Theta_{\varepsilon}})^2 ds\Big] + G^2(\varepsilon)\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0} - \hat{\pi}_s^{\Theta_{\varepsilon}})^2 ds\Big] + G^2(\varepsilon)\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big]\Big).$$

• Because of (3.13) and (vi) in Assumptions 3.1 we notice that

$$\mathbb{E}\Big[\int_0^t (\hat{\pi}_s^{\Theta_0})^2 ds\Big] \le C\Big(\mathbb{E}\Big[\int_0^t (A_s^{\Theta_0})^2 ds\Big] + \mathbb{E}\Big[\int_0^t \int_{\mathbb{R}_0} (B_s^{\Theta_0}(z))^2 \ell(dz) ds\Big]\Big).$$

Using (3.15) and the combination of the above inequalities in (3.16) shows that

$$\begin{split} \mathbb{E}[(\mathcal{L}_{t}^{\Theta_{0}} - \mathcal{L}_{t}^{\Theta_{\varepsilon}})^{2}] &\leq C\Big(\widetilde{G}^{2}(\varepsilon)\Big\{\mathbb{E}\Big[\int_{0}^{t}(A_{s}^{\Theta_{0}})^{2}ds\Big] + \mathbb{E}\Big[\int_{0}^{t}\int_{\mathbb{R}_{0}}(B_{s}^{\Theta_{0}}(z))^{2}\ell(dz)ds\Big]\Big\} \\ &+ \mathbb{E}\Big[\int_{0}^{t}(\bar{A}_{s}^{\varepsilon})^{2}ds\Big] + \mathbb{E}\Big[\int_{0}^{t}\int_{\mathbb{R}_{0}}(\bar{B}_{s}^{\varepsilon}(z))^{2}\ell(dz)ds\Big] + \mathbb{E}\Big[\int_{0}^{t}(C_{s}^{\Theta_{\varepsilon}})^{2}ds\Big] \\ &+ \mathbb{E}\Big[\int_{0}^{t}(\hat{\pi}_{s}^{\Theta_{0}} - \hat{\pi}_{s}^{\Theta_{\varepsilon}})^{2}ds\Big]\Big). \end{split}$$

Finally by Lemma 3.3 and Theorems 3.5 and 3.8 we conclude the proof.

The cost processes of the quadratic hedging strategy for  $\hat{H}_T$ ,  $\hat{H}_T^{\varepsilon}$  are defined by  $K^{\Theta_0} = \mathcal{L}^{\Theta_0} + \hat{V}_0^{\Theta_0}$  and  $K^{\Theta_{\varepsilon}} = \mathcal{L}^{\Theta_{\varepsilon}} + \hat{V}_0^{\Theta_{\varepsilon}}$ . The upcoming result concerns the robustness of the cost process and follows directly from the previous theorem.

Corollary 3.11. Under Assumptions 3.1 and 3.7, there exists a positive constant C such that it holds for all  $t \in [0,T]$  that

$$\mathbb{E}[(K_t^{\Theta_0} - K_t^{\Theta_{\varepsilon}})^2] \le C\Big(\mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2] + \widetilde{G}^2(\varepsilon)\mathbb{E}[\hat{H}_T^2]\Big).$$

*Proof.* Notice that

$$\mathbb{E}[(K_t^{\Theta_0} - K_t^{\Theta_{\varepsilon}})^2] \le 2(\mathbb{E}[(\mathcal{L}_t^{\Theta_0} - \mathcal{L}_t^{\Theta_{\varepsilon}})^2] + \mathbb{E}[(\hat{V}_0^{\Theta_0} - \hat{V}_0^{\Theta_{\varepsilon}})^2]),$$

wherein

$$\mathbb{E}[(\hat{V}_0^{\Theta_0} - \hat{V}_0^{\Theta_\varepsilon})^2] \leq \mathbb{E}\Big[\sup_{0 \leq t \leq T} (\hat{V}_t^{\Theta_0} - \hat{V}_t^{\Theta_\varepsilon})^2\Big].$$

Theorems 3.6 and Theorem 3.10 complete the proof.

The risk processes of the RM strategies are given by

$$(3.17) R_t^{\Theta_0} = \mathbb{E}[(K_T^{\Theta_0} - K_t^{\Theta_0})^2 | \mathcal{F}_t] \quad \text{and} \quad R_t^{\Theta_{\varepsilon}} = \mathbb{E}[(K_T^{\Theta_{\varepsilon}} - K_t^{\Theta_{\varepsilon}})^2 | \widetilde{\mathcal{F}}_t].$$

The following theorem contains the robustness result of the overall risk of the RM strategy at time zero.



**Theorem 3.12.** Let Assumptions 3.1 and 3.7 be in force. For the risk processes  $R^{\Theta_0}$  and  $R^{\Theta_{\varepsilon}}$ , defined in (3.17), we have at time zero that

$$(R_0^{\Theta_0} - R_0^{\Theta_{\varepsilon}})^2 \le C \Big( \mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2] + \widetilde{G}^2(\varepsilon) \mathbb{E}[\hat{H}_T^2] \Big),$$

where C is a positive constant.

*Proof.* By the definitions of the risk and cost process it holds that

$$R_0^{\Theta_0} = \mathbb{E}[(K_T^{\Theta_0} - K_0^{\Theta_0})^2] = \mathbb{E}[(\mathcal{L}_T^{\Theta_0})^2],$$
  

$$R_0^{\Theta_{\varepsilon}} = \mathbb{E}[(K_T^{\Theta_{\varepsilon}} - K_0^{\Theta_{\varepsilon}})^2] = \mathbb{E}[(\mathcal{L}_T^{\Theta_{\varepsilon}})^2].$$

Besides we know that  $\mathcal{L}^{\Theta_0}$  and  $\mathcal{L}^{\Theta_{\varepsilon}}$  are martingales with value zero at time zero. For two random variables X and Y it holds that  $-1 \leq \operatorname{Corr}(X,Y) \leq 1$ . This leads to the inequality

$$|\mathbb{E}[XY]| \le \sqrt{\operatorname{Var}(X)\operatorname{Var}(Y)} + |\mathbb{E}[X]\mathbb{E}[Y]|.$$

Taking the relations above into account we derive

$$\begin{aligned} \left| R_0^{\Theta_0} - R_0^{\Theta_{\varepsilon}} \right| &= \left| \mathbb{E}[(\mathcal{L}_T^{\Theta_0})^2 - (\mathcal{L}_T^{\Theta_{\varepsilon}})^2] \right| = \left| - \mathbb{E}[(\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_{\varepsilon}})^2] + 2\mathbb{E}[\mathcal{L}_T^{\Theta_0} (\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_{\varepsilon}})] \right| \\ &\leq \mathbb{E}[(\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_{\varepsilon}})^2] + 2\sqrt{\operatorname{Var}(\mathcal{L}_T^{\Theta_0}) \operatorname{Var}(\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_{\varepsilon}})} + 2\left| \mathbb{E}[\mathcal{L}_T^{\Theta_0}] \mathbb{E}[\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_{\varepsilon}}] \right|. \end{aligned}$$

Thus we obtain

$$(R_0^{\Theta_0} - R_0^{\Theta_\varepsilon})^2 \le C \Big( (\mathbb{E}[(\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_\varepsilon})^2])^2 + R_0^{\Theta_0} \mathbb{E}[(\mathcal{L}_T^{\Theta_0} - \mathcal{L}_T^{\Theta_\varepsilon})^2] \Big).$$

Since the RM strategy minimises  $R^{\Theta_0}$  we know that  $R_0^{\Theta_0}$  is bounded if there exists an optimal strategy. The proof is fulfilled by using Theorem 3.10.

3.4. Robustness results for the mean-variance hedging. Since the optimal numbers  $\xi^{\Theta_0}$  and  $\xi^{\Theta_{\varepsilon}}$  of risky assets are the same in the RM and the MVH strategy, the amounts of wealth  $\hat{\pi}^{\Theta_0}$  and  $\hat{\pi}^{\Theta_{\varepsilon}}$  and the gain processes  $\hat{G}^{\Theta_0}$  and  $\hat{G}^{\Theta_{\varepsilon}}$  also coincide for both strategies. Therefore we conclude that the robustness results of the amount of wealth and gain process also hold true for the MVH strategy, see Theorems 3.8 and 3.9.

The cost and risk for a MVH strategy are not the same as for the RM strategy. However, under the assumption that a fixed starting amount  $\tilde{V}_0$  is available to set up a MVH strategy, we derive a robustness result for the loss at time of maturity. When the model (2.1) is considered, it holds that the loss at time of maturity T is given by

$$L^{\Theta_0} = \hat{H}_T - \widetilde{V}_0 - \int_0^T \xi_s^{\Theta_0} d\hat{S}_s.$$

On the other hand, for the second model (2.13) the loss at time of maturity T equals

$$L^{\Theta_{\varepsilon}} = \hat{H}_{T}^{\varepsilon} - \widetilde{V}_{0} - \int_{0}^{T} \xi_{s}^{\Theta_{\varepsilon}} d\hat{S}_{s}^{\varepsilon}.$$

When Assumptions 3.1 and 3.7 are imposed, we derive via Theorem 3.9 that

$$\mathbb{E}[(L^{\Theta_0} - L^{\Theta_{\varepsilon}})^2] \le C\Big(\mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2] + \widetilde{G}^2(\varepsilon)\mathbb{E}[\hat{H}_T^2]\Big),$$

for a positive constant C.

Note that we cannot draw any conclusions from the results above about the robustness of the value of the discounted portfolio for the MVH strategy, since the portfolios are strictly different for both strategies.



#### 4. Conclusion

Two different geometric Lévy stock price models were considered in this paper. We proved that the RM and MVH strategy in a martingale setting are stable against the choice of the model. To this end we made use of BSDEJs and the obtained  $L^2$ -convergence rates are expressed in terms of estimates of the form  $\mathbb{E}[(\hat{H}_T - \hat{H}_T^{\varepsilon})^2]$ . The latter estimate is a well studied quantity, see Benth et al. (2011) and Kohatsu-Higa and Tankov (2010). In the current paper, we considered two possible models for the price process. Starting from the initial model (2.1) other models could be constructed by truncating the small jumps and possibly rescaling the original Brownian motion (cfr. Di Nunno et al. (2015)). Similar robustness results hold for quadratic hedging strategies in a martingale setting in these other models.

In Di Nunno et al. (2015) a semimartingale setting was considered and conditions had to be imposed to guarantee the existence of the solutions to the BSDEJs. In this paper however, we observed a martingale setting and there was no driver in the BSDEJs which immediately guaranteed the existence of the solution to the BSDEJs. On the other hand, since the two models were observed under two different martingale measures, we had to fall back on the common historical measure for the robustness study. Therefore, a robustness study of the martingale measures had to be performed and additional terms made some computations more involved compared to the semimartingale setting studied by Di Nunno et al. (2015).

In this approach based on BSDEJs we could not find explicit robustness results for the optimal number of risky assets. Therefore we refer to Daveloose et al. (2014), where a robustness study is performed in a martingale and semimartingale setting based on Fourier transforms. Note that in Daveloose et al. (2014) robustness was mainly studied in  $L^1$ -sense and they noted that their results can be extended into  $L^2$ -convergence, whereas  $L^2$ -robustness results are explicitly derived in the current paper.

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